

Energy-Neutral Source-Channel Coding with Battery and Memory Size Constraints

Paolo Castiglione and Gerald Matz

Abstract—We study energy management policies for the compression and transmission of source data collected by an energy-harvesting sensor node with a finite energy buffer (e.g., rechargeable battery) and a finite data buffer (memory) between source encoder and channel encoder. The sensor node can adapt the source and channel coding rates depending on the observation and channel states. In such a system, the absence of precise information about the amount of energy available in the future is a key challenge. We provide analytical bounds and scaling laws for the average distortion that depend on the size of the energy and data buffers. We furthermore design a resource allocation policy that achieves almost optimal distortion scaling. Our results demonstrate that the energy leakage of state of art energy management policies can be avoided by jointly controlling the source and channel coding rates.

Index Terms—Energy harvesting, source-channel coding, multimedia wireless sensors, optimal power allocation, Markov processes.

I. INTRODUCTION

ENERGY harvesting techniques [1] enable the design of completely autonomous wireless sensor networks (WSN). However, fluctuations in the amount of the energy being harvested call for resource management policies that achieve a trade-off between short-term metrics like delay and data queue length and long-term performance indicators like throughput and average distortion (see [2] and references therein).

In a WSN, an additional challenge is the fact that the energy consumption of source compression is in the same order as that of transmission. Even without energy harvesting this allows for energy savings via a joint energy management for source coding and transmission [3]–[5]. These results have been extended to fluctuating energy sources in [6], [7].

In this paper, we consider a single sensor node and adopt the model from [6], where an *energy buffer*, e.g., a rechargeable battery, stores the harvested energy. In each time slot, the node acquires and compresses an observation with a suitably adapted rate. The observation is characterized by a time-varying state, e.g., observation signal-to-noise ratio (SNR). Then, the node stores the source coder output bits in a *data*

buffer (memory). Furthermore, it transmits to the destination a certain number of bits from the data buffer, using a suitably adapted channel coding rate. The transmission channel is characterized by an instantaneous channel SNR.

In previous work [6], we characterized optimal energy management policies that achieve minimum distortion for the extreme cases where the energy and data buffer are either infinite or very small. For infinite buffer size, where the stability of the data queue needs to be guaranteed, the optimal policies independently allocate energy to the source and channel encoders. On the other hand, for the case of very small buffer size, a joint energy allocation by means of dynamic programming was found to be optimal.

In this paper, we consider finite buffers and use large deviation tools for our analysis that were developed in the seminal work of Tse [8]. Compared to dynamic programming, these tools have the advantage of not suffering from the curse of dimensionality. In [8], only the source coding is taken into account in the sensor policy, which thus amounts to choosing a point on the rate-distortion curve. Neither the problem of maintaining energy-neutrality, nor fluctuations of the available energy, nor optimal resource allocation among source and channel encoder have been addressed in [8].

In this work, we claim that distortion optimality can be achieved via a joint energy management for the source and channel encoders. In particular, we provide analytical bounds on the average distortion achievable with an energy-harvesting sensor, and on the scaling laws of the achievable average distortion with respect to buffer size. We further propose a joint energy management for source and channel encoding that asymptotically achieves the distortion lower bound and scales almost optimally with buffer size. We emphasize that in related work [9], [10] on this topic a joint adaptation of the source code and of the channel code has not been considered since the bit stream entering the data buffer was modeled as exogenous (i.e., uncontrollable).

Other recent contributions [7], [11] for multi-hop systems have shown that a good trade-off between performance and buffer sizes can be found by using Lyapunov optimization techniques that do not require knowledge of the statistics of the system states. In particular, [7] addresses the problem of jointly controlling distributed source coding and data transmission and develops policies that achieve a distortion optimality gap that is inversely proportional to buffer size. In contrast to our work, the optimality of such policies is not discussed in [7].

II. SYSTEM MODEL

We consider a system in which a single energy-harvesting sensor node communicates with a single receiver. A block

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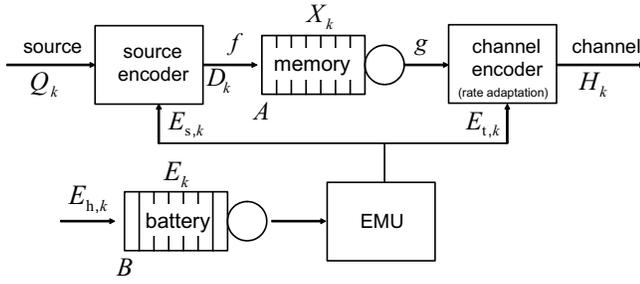


Fig. 1. Model for the energy harvesting sensor.

diagram of the sensor node is depicted in Fig. 1. It essentially consists of a source encoder, a transmitter, an energy buffer, a data buffer, and an energy management unit (EMU).

Energy buffer. In our model, the sensor operation is structured in time slots (indexed by k). The energy harvested in slot k , denoted $E_{h,k} \in \mathbb{R}_+$, is accumulated in an energy buffer of finite size B , hence-forth also referred to as battery. For convenience, all energies are normalized by the number N of channel uses per slot (i.e., the number of symbols transmitted per time slot). The harvested energy $E_{h,k}$ is assumed to be a discrete stationary irreducible aperiodic Markov process. The steady-state probability density function (pdf) of $E_{h,k}$ is denoted $p_E(e)$. The energy E_{k+1} available in the battery for use in slot $k+1$ evolves as

$$E_{k+1} = \min\{B, [E_k - (E_{s,k} + E_{t,k})]^+ + E_{h,k}\}, \quad (1)$$

where $[x]^+ = \max\{0, x\}$. Here, $[E_k - (E_{s,k} + E_{t,k})]^+$ is the residual energy from the previous slot, with $E_{s,k}$ and $E_{t,k}$ denoting the energies allocated in slot k for source encoding and transmission, respectively. We do not take into account the energy consumed by channel encoding and channel state acquisition, since they are typically small compared to the transmit energy in the scenario considered. Energy-neutrality amounts to the constraint $E_{s,k} + E_{t,k} \leq E_k$.

Source encoder. The sensor takes M measurements per time slot. The quality of these measurements is characterized by a parameter sequence $Q_k \in \mathcal{Q}$, which is assumed to form a discrete stationary irreducible aperiodic Markov process. As an example, Q_k could model the measurement SNR, which may change over time due to source movement or environmental factors. The set \mathcal{Q} is assumed to be discrete and finite. The steady-state probability mass function (pmf) for Q_k is denoted $\Pr(q) = \Pr(Q_k = q)$, $q \in \mathcal{Q}$. Due to sampling, analog-to-digital conversion, and compression the sensor acquires the source in a lossy fashion. The loss is captured by the distortion $D_k \in \mathbb{R}^+$, obtained from a given distortion metric such as the mean square error (MSE). The bit stream resulting at the source encoder output is stored in a data buffer, subsequently also referred to as memory.

The number of bits produced by the source encoder within slot k is given by $f^{(q)}(D_k, E_{s,k}) = f(D_k, E_{s,k}, Q_k = q)$. Here, the rate-distortion-energy function f models the dependence of the source encoder output on the distortion level D_k , the allocated energy $E_{s,k}$, and the observation state Q_k . The function $f^{(q)}(D_k, E_{s,k})$ is assumed (for any $q \in \mathcal{Q}$) to be continuous, differentiable, and separately strictly

convex and non-increasing in D_k and $E_{s,k}$. Example rate-distortion-energy functions are provided in [6]. Conventional rate-distortion functions are special cases without dependence on $E_{s,k}$.

Transmitter. The channel between sensor and destination is characterized by a discrete stationary irreducible aperiodic Markov process $H_k \in \mathcal{H}$ that changes slowly over time (e.g., block-fading). The pmf of H_k is given by $\Pr(h) = \Pr(H_k = h)$, $h \in \mathcal{H}$. The transmitter uses the channel N times per slot. A maximum number $g^{(h)}(E_{t,k}) = g(H_k = h, E_{t,k})$ of bits per slot can be communicated successfully to the destination. For any $h \in \mathcal{H}$, the channel rate function $g^{(h)}(E_{t,k})$ is assumed to be continuous, differentiable, strictly concave, and non-decreasing in $E_{t,k}$; furthermore, $g^{(h)}(0) = 0$. We consider rate-adaptive transmission schemes that achieve arbitrarily small block error probabilities. An example for g is given by the Shannon capacity of the additive white Gaussian noise (AWGN) channel with SNR h . However, the channel-rate function can also model the rate of channel codes with a non-zero gap to Shannon capacity. The number of bits actually transmitted using the allocated energy $E_{t,k}$ is given by $\min\{X_k, g^{(h)}(E_{t,k})\}$.

Data buffer. The size and queue length of the data buffer are denoted by A and X_k , respectively. The data queue length evolves as

$$X_{k+1} = \min\{A, [X_k - g(H_k, E_{t,k})]^+ + f(D_k, E_{s,k}, Q_k)\}. \quad (2)$$

The source encoder increases the data queue length by $f(D_k, E_{s,k}, Q_k)$ bits while the transmitter decreases the queue length by transmitting $g(H_k, E_{t,k})$ bits. When all parameters except $E_{s,k}$ and $E_{t,k}$ are fixed, (2) captures the trade-off that results from splitting the available energy between the source encoder and the transmitter. Ideally, it is desirable to decrease $f(D_k, E_{s,k}, Q_k)$ by increasing $E_{s,k}$ and simultaneously increase $g(H_k, E_{t,k})$ by increasing $E_{t,k}$. However, due to energy neutrality $E_{s,k}$ and $E_{t,k}$ cannot be simultaneously increased without bounds.

III. PROBLEM STATEMENT AND MAIN RESULTS

A. Problem Statement

The results obtained in what follows are based on the assumption that the buffer sizes A and B are much larger than the maximum variation of the respective buffer states, i.e., $(A \gg \max |X_{k+1} - X_k|)$ and $(B \gg \max |E_{k+1} - E_k|)$. We further note that the Markov assumption for the energy harvesting, for the observation state, and for the channel state generalizes the memoryless assumption that was used in related work [2], [9], and is inspired by recent models for real harvesting processes [12] as well as by well-established models for the wireless channel [13]. An extension to even more general models is beyond the scope of this paper.

The EMU has to prescribe the distortion D_k and the energies $E_{s,k}$ and $E_{t,k}$ to be allocated to the source encoder and the transmitter, respectively. It does so using the combined state of the energy buffer, the data buffer, the source, and the channel, formally $S_k = \{E_k, X_k, Q_k, H_k\}$. More specifically, the EMU uses a policy $\pi = \{\pi_k\}_{k \geq 1}$

where $\pi_k = \{D_k(S^{(k)}), E_{s,k}(S^{(k)}), E_{t,k}(S^{(k)})\}$ determines the parameters $(D_k, E_{s,k}, E_{t,k})$ in the k th time slot based on the present and past states $S^{(k)} = \{S_1, \dots, S_k\}$.

To decrease the distortion, the sensor can either use more compression energy, thereby faster discharging the energy buffer, or compress less, thereby faster filling the data buffer (which necessitates an increase of the transmission energy to empty the buffer). Therefore, any policy π amounts to a trade-off between the distortion performance and the risk of an energy buffer drain or data buffer overflow. If the battery is empty or the memory is full, a packet is lost and the maximum distortion D_{\max} is accrued. The long-term average distortion achieved with policy π is defined as

$$\bar{D}^\pi = \limsup_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \mathbb{E}^\pi[D_k]. \quad (3)$$

The optimal EMU policy π_{opt} achieves the minimum distortion $\bar{D}^{\min} = \min_\pi \bar{D}^\pi$; hence, $\bar{D}^\pi \geq \bar{D}^{\min}$ for all π .

B. Lower bound on the achievable distortion

We define, conditional on the observation state $Q_k = q$, the long-term average source encoder energy $E_s^{(q)} = \liminf_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \mathbb{E}[E_{s,k}|Q_k = q]$ and the long-term average distortion $D^{(q)} = \liminf_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \mathbb{E}[D_k|Q_k = q]$. Furthermore, the long-term average transmit energy conditional on the channel state $H_k = h$ is defined as $E_t^{(h)} = \liminf_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \mathbb{E}[E_{t,k}|H_k = h]$. Our main results are based on the following convex optimization problem.

Definition 1. The convex optimization problem $\text{CP}(\delta_d, \delta_e)$ is defined as follows:

$$\min_{D^{(q)}, E_s^{(q)}, E_t^{(h)}, \alpha} \sum_q \Pr(q) D^{(q)}$$

subject to

$$\sum_q \Pr(q) f^{(q)}(D^{(q)}, E_s^{(q)}) - \sum_h \Pr(h) g^{(h)}(E_t^{(h)}) \leq \delta_d, \quad (4)$$

$$\sum_q \Pr(q) E_s^{(q)} \leq (1-\alpha)(\mathbb{E}[E_{h,k}] + \delta_e), \quad (5)$$

$$\sum_h \Pr(h) E_t^{(h)} \leq \alpha(\mathbb{E}[E_{h,k}] + \delta_e), \quad (6)$$

$$D^{(q)} \geq 0, \quad E_s^{(q)} \geq 0, \quad E_t^{(h)} \geq 0, \quad 0 < \alpha < 1.$$

Here, δ_d and δ_e can be interpreted as the incremental and decremental drifts¹ for the data buffer and for the energy buffer, respectively. The problem $\text{CP}(\delta_d, \delta_e)$ minimizes the inferior limit of the long-term expected distortion, thereby identifying the values of the above defined long-term expectations $D^{(q)}$, $E_s^{(q)}$, $E_t^{(h)}$ and of the associated parameter α . The latter parameter denotes the ratio between the long-term expected energy spent for transmission and the long-term overall expected energy spent for transmission and source

coding. For the problem $\text{CP}(\delta_d, \delta_e)$ with $\delta_d \leq 0$ and $\delta_e \leq 0$, condition (4) is necessary for the *mean rate stability* of the data queue [14], and conditions (5)-(6) are necessary to meet the energy neutrality requirement. We note that the problems $\text{CP}(\delta_d, \delta_e)$ with $\delta_d > 0$ and $\delta_e > 0$ can be viewed as relaxations of $\text{CP}(0, 0)$.

Using the results in [6], we establish a lower bound on the achievable long-term distortion. The proofs of this and the subsequent results are provided in the Appendix.

Proposition 1. Let \bar{D}^* denote the minimum of the problem $\text{CP}(0, 0)$ (i.e., with zero drift, $\delta_d = 0$, $\delta_e = 0$). Then, the minimum distortion is lower bounded as $\bar{D}^{\min} \geq \bar{D}^*$.

We note that it suffices to prove this result for the special case of infinite data and energy buffer, i.e., $A = \infty$ and $B = \infty$. This also establishes the bound for finite buffer sizes since the assumption $A < \infty$ and $B < \infty$ is more restrictive (an infinite buffer can always mimic a finite buffer) and hence cannot lead to a smaller achievable distortion. The proof substantially demonstrates by means of Jensen inequality that condition (4) is necessary to meet the *mean rate stability* of the infinite data queue [14].

The next result provides a lower bound on the scaling behavior of the difference $\bar{D}^{\min} - \bar{D}^*$, showing how \bar{D}^{\min} converges to \bar{D}^* when buffer size increases.²

Proposition 2. For any EMU policy π there is

$$\bar{D}^{\min} - \bar{D}^* = \Omega(A^{-2}) + \Omega(B^{-2}).$$

This results states that the optimality gap $\bar{D}^{\min} - \bar{D}^*$ is asymptotically bounded below by $c_1 A^{-2} + c_2 B^{-2}$ (here, c_1 and c_2 are constants). Thus, with increasing buffer size, \bar{D}^{\min} cannot converge to the minimum distortion \bar{D}^* at a rate faster than $c_1 A^{-2} + c_2 B^{-2}$. This result is not intuitive. The key idea behind the proof provided in the Appendix entails the manipulation of appropriate balance equations for each buffer as done in [8, 2.4.1].

C. Distortion achievable with finite buffer size

We now present a stationary EMU policy $\pi^0 = \{\pi_k^0\}_{k \geq 1}$ that only depends (in a deterministic manner) on the current state S_k and performs close to the lower bound established in Proposition 2. This policy enforces drifts depending on hyper states that indicate whether the queues are more or less than half full. These hyper states are captured by the indices $n = \mathbb{I}\{X_k \geq A/2\}$ and $m = \mathbb{I}\{E_k < B/2\}$.³ The data queue drift and the energy buffer drift then equal $\delta_d = (-1)^n \frac{\beta_1 \ln A}{A}$ and $\delta_e = (-1)^m \frac{\beta_2 \ln B}{B}$, respectively, with β_1 and β_2 sufficiently large constants (see Appendix). The sign of these drifts ensures that the buffer states are pushed towards the respective center levels $A/2$ and $B/2$. Furthermore, the drift magnitude depends on the size of the respective buffer. For example, the data queue drift decreases with increasing buffer size A . This is

¹The drift δ_d is the long-term expected difference between the size of the data buffer *input* and the size of the data buffer *output*. Vice versa, the drift δ_e is the long-term expected difference between the size of the energy buffer *output* and the size of the energy buffer *input*. For the definition of drift, the buffer is assumed to be unbounded. See Appendix for a more formal definition.

²The following notations are used to compare two sequences a_n and b_n as n grows: $a_n = O(b_n)$ if $a_n/b_n < c$ for all n and some constant c ; $a_n = \Omega(b_n)$ if $b_n = O(a_n)$; $a_n = \Theta(b_n)$ if $a_n = O(b_n)$ and $b_n = O(a_n)$; $a_n = o(b_n)$ if $\lim_{n \rightarrow \infty} a_n/b_n = 0$.

³The indicator function $\mathbb{I}\{\cdot\}$ equals 1 if the argument is true and 0 otherwise.

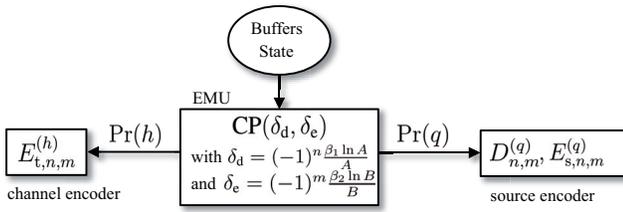


Fig. 2. Illustration of the EMU policy (7).

intuitive since smaller buffers tend to become full faster and hence require a stronger drift to avoid overflow. The same reasoning applies to the battery drift.

Definition 2. We define the policy $\pi_k^0 = \{D_k, E_{s,k}, E_{t,k}\}$ by

$$\begin{aligned} D_k &= D_{n,m}^{(q)}, \text{ for } Q_k = q, \\ E_{s,k} &= \min\{(1 - \alpha_{n,m})E_k, E_{s,n,m}^{(q)}\}, \text{ for } Q_k = q, \\ E_{t,k} &= \min\{\alpha_{n,m}E_k, E_{t,n,m}^{(h)}\}, \text{ for } H_k = h, \end{aligned} \quad (7)$$

where the parameters $D_{n,m}^{(q)}$, $E_{s,n,m}^{(q)}$, $E_{t,n,m}^{(h)}$, and $\alpha_{n,m}$ are obtained by solving the optimization problem $\text{CP}(\delta_d, \delta_e)$ from Definition 1 with the drifts chosen as $\delta_d = (-1)^n \frac{\beta_1 \ln A}{A}$ and $\delta_e = (-1)^m \frac{\beta_2 \ln B}{B}$.

For given rate functions f and g and given statistics for Q_k and H_k , the above problem must be solved for all four possible hyper states using standard convex optimization tools and the resulting parameters of the policy can then be stored in a lookup table. As illustrated in Fig. 2, the source code is determined by the distortion $D_{n,m}^{(q)}$ and the energy consumption $E_{s,n,m}^{(q)}$ depending on the state of the source q . On the other hand, the channel code is determined by the energy consumption $E_{t,n,m}^{(h)}$ depending on the channel state h . If the energy in the battery is not sufficient to provide $E_{s,n,m}^{(q)}$ and $E_{t,n,m}^{(h)}$, policy (7) assigns the residual energy E_k to the source encoder and to the channel encoder according to parameter $\alpha_{n,m}$. The next result assesses the performance of the EMU policy defined above.

Proposition 3. *The policy π^0 achieves a long-term average distortion \bar{D}^{π^0} that approaches \bar{D}^* as $O(A^{-2} \ln^2 A) + O(B^{-2} \ln^2 B)$.*

Proposition 3 states that the scaling behavior of the long-term average distortion achieved with the policy π^0 is almost optimal (cf. Proposition 2). More specifically, the optimality gap $\bar{D}^{\pi^0} - \bar{D}^*$ converges to zero as $O(A^{-2} \ln^2 A) + O(B^{-2} \ln^2 B)$. This scaling behaviour can be interpreted as the truncation of the Taylor representation of the optimality gap to the second order derivative with respect to the drifts in each buffer. The second order component dominates the performance because the first order component cancels out due to the fact that the drifts in each buffer have opposite signs, are equal in magnitude and occur with asymptotically equal probability (see Appendix). Moreover, the data and energy buffer drifts imposed by π^0 keep the probabilities of battery depletion and memory overflow small (but different from zero), such that they become asymptotically negligible. More complex EMU policies, for instance, that can online

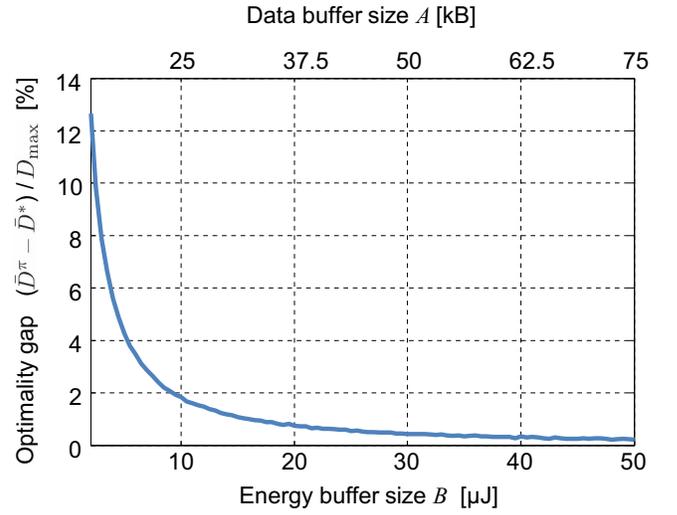


Fig. 3. Optimality gap versus data buffer size and energy buffer size.

adapt the source-channel code and the associated energies, might force these probabilities to zero. However, according to Proposition 2, any other policy, even if more complex and adaptive, cannot perform substantially better than π^0 .

We additionally observe that the source and channel encoder parameters are jointly adapted over time. This is consistent with the dynamic programming solution in [6] for small memory/battery sizes. It is also interesting to note that the source encoder and the transmitter are *separately* controlled as long as the hyper state of the buffers remains the same, for instance, as long as the data queue length is less than $A/2$ and the available energy is larger than $B/2$. Hence, as $A, B \rightarrow \infty$, a separate energy management for source encoding and transmission remains optimal, which is consistent with the results in [6].

IV. NUMERICAL RESULTS

We present Monte Carlo simulations in order to numerically assess the performance of the proposed EMU policy. We consider a system with slot duration 10 ms. In each, the sensor acquires $M = 10^3$ noisy samples with SNR Q_k . The source encoder output bits are passed through a channel encoder and transmitted over an AWGN channel with $N = 2 \cdot 10^3$ channel uses (the transmission bandwidth thus is 200kHz).

Using the model in [3], the source encoder is characterized by the rate-distortion-energy function $f^{(q)}(D_k, E_{s,k}) = M \log_2 \left(\frac{D_{\max} - D_{\text{mmse}}}{D_k - D_{\text{mmse}}} \right) \xi(E_{s,k})$. Here, the first term is the information-theoretic rate-distortion limit for a zero-mean white Gaussian source with variance D_{\max} and with minimum distortion (minimum MSE) $D_{\text{mmse}} = (q + 1/D_{\max})^{-1}$. The function $\xi(E_{s,k}) = \zeta \max \left\{ 1, \left(\frac{N}{M} \frac{E_{s,k}}{E_{s,\max}} \right)^\eta \right\}$ accounts for the rate increase incurred by practical (energy-limited) compression schemes. For our simulations we chose $\zeta = 2$, $\eta = -2/3$, and a maximum energy consumption per slot of $\frac{M}{N} E_{s,\max} = 1 \text{ nJ} = -7 \text{ dBm}$. The transmitter is characterized by the channel rate function $g^{(h)}(E_{t,k}) = N \log \left(1 + \frac{E_{t,k}}{h\Gamma\sigma} \right)$, where $H_k = h$ accounts for the path loss, $\Gamma = 7 \text{ dB}$ is the SNR gap to Shannon capacity of the (rate-adaptive) channel code, and σ is the noise power (measured in J, here -80 dBm).

The source SNR $\mathcal{Q} = \{1, 10\}$ dB and the path loss $\mathcal{H} = \{40, 50\}$ dB are two-state Markov chains with uniform steady-state pmf and transitions from one state to the other happening with probability 0.1. The harvested energy $E_{h,k}$ is a uniformly distributed Markov chain with nine states, uniformly spaced in the interval $(0, 100)\mu\text{J/s}$ and with transition probability from each state to any of the other eight states equal to 0.05.

Fig. 3 shows the optimality gap of the policy π^0 versus data and energy buffer size. It is seen that the optimality gap can be decreased to 0.2% of the source variance D_{\max} by simultaneously increasing the energy buffer size and the data buffer size to realistic values of 50 μJ and 75 kB, respectively.⁴ Moreover, these results confirm the validity of the scaling behavior stated in Proposition 3.

V. CONCLUDING REMARKS

In this paper we have proposed an energy management policy for energy-harvesting sensor nodes that achieves a close-to-optimal distortion scaling with respect to battery and memory size. Our large deviations results substantially differ from [9], [10], which assumed the bits arriving in the data buffer to be exogenous (uncontrolled). In these papers, the average harvested energy was assumed to be strictly larger than the average energy required to achieve the optimal utility. The energy buffer is therefore constantly filled, which implies that only the data buffer needs to be controlled, at the price that not all harvested energy is used. In contrast, our proposed energy management policy jointly adapts the source code and the channel code, which leads to a separate control of the energy buffer and the data buffer and achieves the distortion lower bound without any detrimental energy leakage.

APPENDIX

Proof of Proposition 1

We prove this result for the special case of infinite data and energy buffer, i.e., $A = \infty$ and $B = \infty$. This also establishes the bound for finite buffer sizes since the assumption $A < \infty$ and $B < \infty$ is more restrictive (an infinite buffer can always mimic a finite buffer) and hence cannot lead to a smaller achievable distortion.

Without loss of generality we assume *mean rate stability*, i.e., $\limsup_{n \rightarrow \infty} \frac{1}{n} \mathbb{E}[X_k] = 0$ and $\mathbb{E}[X_0] < \infty$, both of which are always satisfied for the case of finite data buffers. Using [14, Theorem 3], we now prove that mean rate stability implies the same necessary stability conditions as in [6, Proposition 1].

Since f is convex non-increasing in D_k and $E_{s,k}$ and g is concave non-decreasing in $E_{t,k}$, Jensen's inequality implies

$$\begin{aligned} & \frac{1}{n} \sum_{k=1}^n \mathbb{E}[f(D_k, E_{s,k}, Q_k) - g(H_k, E_{t,k})] \\ & \geq \sum_q \Pr(q) f^{(q)}(\bar{D}_n^{(q)}, \bar{E}_{s,n}^{(q)}) - \sum_h \Pr(h) g^{(h)}(\bar{E}_{t,n}^{(h)}). \end{aligned}$$

⁴These are typical values for the capacitor and for the memory of a low-power sensor node.

with

$$\begin{aligned} \bar{D}_n^{(q)} &= \frac{1}{n} \sum_{k=1}^n \mathbb{E}[D_k | Q_k = q], \\ \bar{E}_{s,n}^{(q)} &= \frac{1}{n} \sum_{k=1}^n \mathbb{E}[E_{s,k} | Q_k = q], \\ \bar{E}_{t,n}^{(h)} &= \frac{1}{n} \sum_{k=1}^n \mathbb{E}[E_{t,k} | H_k = h]. \end{aligned}$$

Furthermore, we have

$$\begin{aligned} \limsup_{n \rightarrow \infty} \sum_q \Pr(q) f^{(q)}(\bar{D}_n^{(q)}, \bar{E}_{s,n}^{(q)}) &= \sum_q \Pr(q) f^{(q)}(D^{(q)}, E_s^{(q)}), \\ \limsup_{n \rightarrow \infty} - \sum_h \Pr(h) g^{(h)}(\bar{E}_{t,n}^{(h)}) &= - \sum_h \Pr(h) g^{(h)}(E_t^{(h)}), \end{aligned}$$

with $D^{(q)} = \liminf_{n \rightarrow \infty} \bar{D}_n^{(q)}$, $E_s^{(q)} = \liminf_{n \rightarrow \infty} \bar{E}_{s,n}^{(q)}$, and $E_t^{(h)} = \liminf_{n \rightarrow \infty} \bar{E}_{t,n}^{(h)}$. Using the fact that mean rate stability and [14, (12)] imply $\limsup_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \mathbb{E}[f(D_k, E_{s,k}, Q_k) - g(H_k, E_{t,k})] \leq 0$, we finally arrive at

$$\sum_q \Pr(q) f^{(q)}(D^{(q)}, E_s^{(q)}) \leq \sum_h \Pr(h) g^{(h)}(E_t^{(h)}),$$

which is the same necessary stability condition as in [6, Proposition 1]. The proof of Proposition 1 thus follows from [6, Proposition 1].

Proof of Proposition 2

We first define some quantities that are instrumental for the proofs of Propositions 2 and 3. The probabilities that a policy π results in an empty energy buffer or a full data queue are respectively defined as

$$p_{\text{EB}}^\pi = \limsup_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \Pr([E_k - E_{s,k} - E_{t,k}]^+ = 0), \quad (8)$$

$$p_{\text{FQ}}^\pi = \limsup_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \Pr(X_k = A) \quad (9)$$

The *decremental* drift δ_e of the unbounded energy buffer queue process $\tilde{E}_{k+1} = \tilde{E}_k - E_{s,k} - E_{t,k} + E_{h,k}$, and the *incremental* drift δ_d of the unbounded data queue process $\tilde{X}_{k+1} = \tilde{X}_k + f(D_k, E_{s,k}, Q_k) - g(H_k, E_{t,k})$ are respectively defined as

$$\delta_e = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \mathbb{E}[\tilde{E}_k - \tilde{E}_{k+1}],$$

$$\delta_d = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \mathbb{E}[\tilde{X}_{k+1} - \tilde{X}_k].$$

The drift δ_d can be viewed as the long-term expected difference between the size of the data buffer *input* and the size of the data buffer *output*, whereas the drift δ_e can be viewed as the long-term expected difference between the size of the energy buffer *output* and the size of the energy buffer *input*.

We denote the event of *normal operation* by $\mathcal{E}_k = \{X_k < A \text{ and } [E_k - (E_{s,k} + E_{t,k})]^+ > 0\}$ (i.e., the data buffer is not full and the energy buffer is not empty). The complementary event is denoted as $\bar{\mathcal{E}}_k$ (i.e., either the data

buffer is full or the energy buffer is empty). The instantaneous expected distortion can now be written as $\mathbb{E}^\pi[D_k] = \mathbb{E}^\pi[D_k|\mathcal{E}_k] \Pr(\mathcal{E}_k) + \mathbb{E}^\pi[D_k|\bar{\mathcal{E}}_k] \Pr(\bar{\mathcal{E}}_k)$. With $\Pr(\mathcal{E}_k) \leq 1$, $\mathbb{E}^\pi[D_k|\bar{\mathcal{E}}_k] = D_{\max}$, and the union bound $\Pr(\bar{\mathcal{E}}_k) \leq \Pr(X_k = A) + \Pr([E_k - (E_{s,k} + E_{t,k})]^+ = 0)$, we obtain the following upper bound on the long-term average distortion \bar{D}^π :

$$\bar{D}^\pi = \limsup_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \mathbb{E}^\pi[D_k] \leq D^{\text{op}} + D_{\max} (p_{\text{FQ}}^\pi + p_{\text{EB}}^\pi). \quad (10)$$

Here, we have used (8), (9), and the long-term average distortion during normal operation,

$$D^{\text{op}} = \limsup_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \mathbb{E}^\pi[D_k|\mathcal{E}_k]. \quad (11)$$

Note that here we assume that if the energy buffer is empty or the data buffer is full, a packet is lost and the maximum distortion D_{\max} is accrued (i.e., the decoder treats this missing packet as an arbitrary vector, e.g., the mean of the source distribution). An energy buffer discharge or data buffer overflow could be handled in a more sophisticated manner, but this does not bear on our asymptotic analysis (see also [8] for a similar reasoning). The bound (10) indicates that studying the optimal convergence of \bar{D}^π to the lower bound \bar{D}^* (see Proposition 1) is equivalent to finding the optimal scaling laws for i) the probabilities p_{EB}^π and p_{FQ}^π approaching zero and ii) the operational distortion D^{op} approaching \bar{D}^* .

In order to prove Proposition 2, we need the following two lemmas.

Lemma 1. *Let $B = \infty$ and consider an arbitrary control scheme $\pi_k = \{D_k, E_{s,k}, E_{t,k}\}$ that achieves $p_{\text{FQ}}^\pi = o(1/A^2)$. Then, $\bar{D}^\pi - \bar{D}^* = \Omega(1/A^2)$.*

Lemma 1 states that for an infinitely large energy buffer no control scheme can make both p_{FQ}^π and \bar{D}^π converge at a rate faster than $1/A^2$. The proof of Lemma 1 is based on [8, Proposition 2.4.1]. Let us consider the optimization problem $\text{CP}(\delta_d, 0)$, i.e., the problem formulated in Def. 1 with the specific choice $\delta_e = 0$. Denote the minimum of $\text{CP}(\delta_d, 0)$ by $D_T(\delta_d)$.

According to Proposition 1, the solution to this problem with $\delta_d = 0$ equals the lower bound on the minimum achievable distortion. Notice that, by the convexity [15] of the problem, $D_T(\delta_d)$ is convex and non-increasing in δ_d . Moreover, using the same arguments as in Proposition 1, it can be proved that there exists no policy π that is able to achieve a long term average distortion smaller than $D_T(\delta_d)$ with a data queue drift smaller or equal to δ_d . Thus, $D_T(\delta_d)$ can be viewed as the lower limit of the distortion-drift region. This observation allows us to directly apply the proof of [8, Proposition 2.4.1] to Lemma 1.

Lemma 2. *Let $A = \infty$, and consider an arbitrary control scheme $\pi_k = \{D_k, E_{s,k}, E_{t,k}\}$ that achieves $p_{\text{EB}}^\pi = o(1/B^2)$. Then, $\bar{D}^\pi - \bar{D}^* = \Omega(1/B^2)$.*

Lemma 2 is the counterpart of Lemma 1; it states that, for an infinitely large data buffer size, no control scheme can

achieve a convergence rate faster than $1/B^2$ for both p_{EB}^π and \bar{D}^π . The proof parallels that of Lemma 1.

We now prove Proposition 2 by contradiction. Assume that there exists a policy with $p_{\text{EB}}^\pi = o(1/B^2)$ and $p_{\text{FQ}}^\pi = o(1/A^2)$ that achieves $\bar{D}^\pi - \bar{D}^* = O(1/A^2) + O(1/B^2)$, i.e., $\bar{D}^\pi - \bar{D}^*$ is asymptotically bounded above by $c_1/A^2 + c_2/B^2$ (where c_1 and c_2 are constant factors). Such a policy would violate Lemma 1 (or Lemma 2) as B (or A) tends to infinity and hence cannot exist. This implies that $\bar{D}^\pi - \bar{D}^* = \Omega(1/A^2) + \Omega(1/B^2)$, which concludes the proof of Proposition 2.

Proof of Proposition 3

In order to prove Proposition 3, we first need to recall some known results. Let us define the random walk $Z_{k+1} = Z_k + W_k$, $k \geq 0$, with $Z_0 = 0$ and W_k a stationary, irreducible, and aperiodic Markov chain with states w_i , $1 \leq i \leq I$. The transition probability from state w_i to state w_j is denoted $p_{i,j}$ and $\Pr(w_i)$ is the invariant distribution of W_k . The drift $R = \mathbb{E}[W_k]$ of the random walk is assumed negative. The moment generating function of W_k is given by

$$\rho(r) = \mathbb{E}[\exp(rW_k)] = \sum_{i=1}^I \Pr(w_i) \exp(rw_i).$$

It can be shown that the function $\log \rho(r)$ (i.e., the cumulant-generating function) has a unique positive zero at $r = r^*$.

Theorem 1. (Wald's identity) *Let K be the first $k \geq 1$ for which $Z_k \geq a \geq 0$ or $Z_k \leq b \leq 0$. Then*

$$\mathbb{E}[\exp(r^* Z_K)] = 1,$$

where r^* is the unique positive root of $\log \rho(r)$. Furthermore,

$$\mathbb{E}[K] \mathbb{E}[W_k] = \mathbb{E}[Z_K].$$

Using Theorem 1 we can compute the probability $p = \Pr(Z_K \geq L)$ that a negative-drift random walk that starts at zero will cross the barrier $L > 0$ before returning to the origin,

$$p \mathbb{E}[\exp(r^* Z_K) | Z_K \geq L] + (1-p) \mathbb{E}[\exp(r^* Z_K) | Z_K \leq 0] = 1.$$

Since $\mathbb{E}[\exp(r^* Z_K) | Z_K \geq L] = \Theta(\exp(r^* L))$ and $\mathbb{E}[\exp(r^* Z_K) | Z_K \leq 0] = \Theta(1)$ in the regime of large L , we have

$$p = \Theta(\exp(-r^* L)),$$

$$\mathbb{E}[Z_K] = p \mathbb{E}[Z_K | Z_K \geq L] + (1-p) \mathbb{E}[Z_K | Z_K \leq 0] = \Theta(1).$$

Hence, by Theorem 1, also the expected crossing time is dominated by the return to the origin, i.e., $\mathbb{E}[K] = \mathbb{E}[Z_K] / \mathbb{E}[W_k] = \Theta(1)$.

Let us now define $D_T(\delta_d, \delta_e)$, as the minimum distortion in the convex problem $\text{CP}(\delta_d, \delta_e)$ (Def. 1). According to Proposition 1, $\bar{D}^* = D_T(0, 0)$ is a lower bound for the minimum achievable distortion $D_T(\delta_d, \delta_e)$. The convexity of $\text{CP}(\delta_d, \delta_e)$ implies that $D_T(\delta_d, \delta_e)$ is convex and non-increasing in (δ_d, δ_e) . Moreover, it can be proved using the same arguments as in Proposition 1 that there exists no policy π that is able to achieve a long-term average distortion \bar{D}^π smaller than $D_T(\delta_d, \delta_e)$ with a data queue drift and energy buffer drift less than or equal to δ_d and δ_e , respectively. Thus,

$D_T(\delta_d, \delta_e)$ can be viewed as a lower bound for the distortion-drift region.

Consider now the policy π° in Def. 2 and recall that the hyper states $\hat{X}_k = \mathbb{I}(X_k \geq A/2) \in \{0, 1\}$ and $\hat{E}_k = \mathbb{I}(E_k < B/2) \in \{0, 1\}$ indicate, respectively, whether the data buffer is more than half full and the energy buffer is more than half empty. Both \hat{X}_k and \hat{E}_k can be shown to be irreducible and aperiodic Markov chains. Furthermore, the data queue *increment* process $W_k^d(\hat{X}_k) = f(D_k, E_{s,k}, Q_k) - g(H_k, E_{t,k})$ in each half of the data buffer is function of the aggregate Markov chain (\hat{E}_k, Q_k, H_k) . Hence, $W_k^d(\hat{X}_k)$ in each half of the data queue is itself an irreducible and aperiodic Markov chain whose mean equals the drift δ_d . Similarly, the energy buffer *decrement* process $W_k^e(\hat{E}_k) = E_{s,k} + E_{t,k} - E_{h,k}$ in each half of the energy buffer is a function of the aggregate Markov chain $(\hat{X}_k, E_{h,k}, Q_k, H_k)$ and hence is itself an irreducible and aperiodic Markov chain whose mean equals the drift (as defined above), i.e., δ_e . We next state a lemma that follows from [10] and relates the unique positive roots $r_d^*(\hat{X}_k)$ and $r_e^*(\hat{E}_k)$ of the cumulant-generating functions of $W_k^d(\hat{X}_k)$ and $W_k^e(\hat{E}_k)$, respectively.

Lemma 3. *Assume that the policy π° is the unique optimal policy for the drifts δ_d and δ_e . Then*

$$\left. \frac{dr_d^*(\hat{X}_k)}{d\delta_d} \right|_{\delta_d=\delta_e=0} = -\frac{2}{\text{var}(W_k^d(\hat{X}_k)|\delta_d=\delta_e=0)}, \quad (12a)$$

$$\left. \frac{dr_e^*(\hat{E}_k)}{d(\delta_e)} \right|_{\delta_d=\delta_e=0} = -\frac{2}{\text{var}(W_k^e(\hat{E}_k)|\delta_d=\delta_e=0)}. \quad (12b)$$

The proof of Lemma 3 is based on the fact that the policy π° depends smoothly on the drifts (δ_d, δ_e) in the neighborhood of $\delta_d = \delta_e = 0$. This implies that $r_d^*(\hat{X}_k)$ and $r_e^*(\hat{E}_k)$ are smooth functions of the means of the respective increment/decrement processes, i.e., δ_d and δ_e , too. Since the supports of $W_k^d(\hat{X}_k)$ and $W_k^e(\hat{E}_k)$ are finite and the policy π° is almost always continuous and differentiable (like the waterfilling-like policy in [16]), the respective second-order derivatives around $\delta_d = \delta_e = 0$ almost always exist and are continuous.

To obtain r_d^* (we omit the argument \hat{X}_k in what follows for notational simplicity), we need to find the root of the cumulant generating function of $W_k^d(\hat{X}_k)$ with $\delta_e = 0$:

$$\Lambda(r_d^*) = \log \rho(r_d^*) = \log \left(\sum_{i=1}^I \Pr(w_i^d) \exp(r_d^* w_i^d) \right) = 0;$$

here, $\Pr(w_i^d)$ is the invariant distribution of $W_k^d(\hat{X}_k)$, which depends on δ_d . Denoting by κ_n the n th cumulant of $W_k^d(\hat{X}_k)$ (i.e., the n th derivative of the cumulant generating function $\Lambda(r)$ at $r = 0$), the Maclaurin expansion of $\Lambda(r_d^*)$ reads

$$\Lambda(r_d^*) = \kappa_0 + \sum_{n=1}^{\infty} \kappa_n \frac{(r_d^*)^n}{n!} = \delta_d r_d^* + \sum_{n=2}^{\infty} \kappa_n \frac{(r_d^*)^n}{n!},$$

where the second equality is obtained with $\kappa_0 = 0$ and $\kappa_1 = \delta_d$. Setting this expression equal to zero, dividing both sides by r_d^* and differentiating with respect to δ_d yields

$$\sum_{n=2}^{\infty} \kappa_n \frac{(n-1)(r_d^*)^{n-2}}{n!} \frac{dr_d^*}{d\delta_d} = -1.$$

Since $r_d^* \rightarrow 0$ as $\delta_d \rightarrow 0$ (due to well-known properties of the moment generating function, see, e.g., [8]), the above expression becomes

$$\left. \frac{\kappa_2}{2} \frac{dr_d^*}{d\delta_d} \right|_{\delta_d=0} = -1.$$

By substituting $\kappa_2 = \text{var}(W_k^d(\hat{X}_k)|\delta_d = \delta_e = 0)$, we obtain (12a) in Lemma 3. The proof of (12b) is analogous.

Let us next consider the probability of a full data buffer $p_{\text{FQ}}^{\pi^\circ}$ obtained by adopting the policy π° . The times at which the data buffer becomes full can be viewed as the epochs of a renewal process [17]. The fullness probability can thus be written as $p_{\text{FQ}}^{\pi^\circ} = 1/\mathbb{E}[Y]$, where Y is the duration between successive time instants at which the buffer becomes full. We define the random walk $Z_k^u = Z_{k-1}^u + W_k^d(\hat{X}_k = 1)$, $k = 0, 1, \dots$, in the upper half of the data buffer, with negative drift δ_d . Let K^u be the smallest time index $k \geq 1$ for which $Z_k^u \geq 0$ or $Z_k^u \leq -A/2$. We then have (for similar arguments refer to [8])

$$\begin{aligned} \mathbb{E}[Y] &= \mathbb{E}[K^u | Z_0^u = 0] + \Pr \left[Z_{K^u}^u \leq -\frac{A}{2} \mid Z_0^u = 0 \right] \\ &\quad \cdot \int_0^{g^{\max}} \Pr \left[Z_{K^u}^u = -\frac{A}{2} - x \mid Z_{K^u}^u \leq -\frac{A}{2}, Z_0^u = 0 \right] \\ &\quad \cdot \mathbb{E} \left[V \mid X_0 = \frac{A}{2} - x \right] dx, \end{aligned} \quad (13)$$

where g^{\max} is the maximum undershoot relative to the center of the buffer (i.e., the maximum variation for the considered policy towards the empty buffer) and V is the time it takes to fill up the buffer starting from an initial buffer state X_0 . We furthermore define the time J^u to be the smallest $k \geq 1$ for which $Z_k^u \leq 0$ or $Z_k^u \geq A/2$. Also, we define another random walk for the dynamics in the lower half of the buffer (state $\hat{X}_k = 0$), $Z_k^l = Z_{k-1}^l + W_k^d(\hat{X}_k = 0)$, $k = 0, 1, \dots$, with positive drift δ_d . The smallest $k \geq 1$ such that $Z_k^l \leq 0$ or $Z_k^l \geq A/2$ is denoted K^l , and the smallest $k \geq 1$ such that $Z_k^l \geq 0$ or $Z_k^l \leq -A/2$ is denoted J^l . For $0 \leq x < g^{\max}$ we have

$$\begin{aligned} \mathbb{E} \left[V \mid X_0 = \frac{A}{2} - x \right] &= \mathbb{E} \left[P \mid X_0 = \frac{A}{2} - x \right] \\ &\quad + \int_0^{f^{\max}} \Pr \left[X_P = \frac{A}{2} + y \mid X_0 \leq \frac{A}{2} - x \right] \\ &\quad \cdot \mathbb{E} \left[V \mid X_0 = \frac{A}{2} + y \right] dy, \end{aligned} \quad (14)$$

where the integral is up to the maximum overshoot f^{\max} relative to the center of the buffer (i.e. the maximum variation for the considered policy towards the full buffer). The quantity P is the first time instant for which the buffer becomes more than half full starting from an initial buffer state X_0 .

By conditioning on the event that the buffer becomes full or that the data queue leaves the upper buffer half, we have for every $0 \leq y < f^{\max}$

$$\begin{aligned} \mathbb{E} \left[V \mid X_0 = \frac{A}{2} + y \right] &= \mathbb{E} \left[J^u \mid Z_0^u = y \right] \\ &\quad + \Pr \left[Z_{J^u}^u \leq 0 \mid Z_0^u = y \right] \int_0^{g^{\max}} \Pr \left[Z_{J^u}^u = -w \mid Z_{J^u}^u \leq 0 \right] \end{aligned}$$

$$\cdot \mathbb{E} \left[V \left| X_0 = \frac{A}{2} - w \right. \right] dw. \quad (15)$$

From (15) it follows that for every $0 \leq y < f^{\max}$

$$\begin{aligned} \mathbb{E} \left[V \left| X_0 = \frac{A}{2} + y \right. \right] &\leq \mathbb{E} \left[J^u \left| Z_0^u = y \right. \right] \\ &+ \Pr[Z_{J^u}^u \leq 0 | Z_0^u = y] \max_{0 \leq w < g^{\max}} \mathbb{E} \left[V \left| X_0 = \frac{A}{2} - w \right. \right]. \end{aligned}$$

Inserting this bound into (14) yields for every $0 \leq x < g^{\max}$

$$\begin{aligned} \mathbb{E} \left[V \left| X_0 = \frac{A}{2} - x \right. \right] &\leq \mathbb{E} \left[P \left| X_0 = \frac{A}{2} - x \right. \right] \\ &+ \max_{0 \leq y < f^{\max}} \mathbb{E} [J^u | Z_0^u = y] \\ &+ \max_{0 \leq y < f^{\max}} \Pr[Z_{J^u}^u \leq 0 | Z_0^u = y] \\ &\cdot \max_{0 \leq w < g^{\max}} \mathbb{E} \left[V \left| X_0 = \frac{A}{2} - w \right. \right]. \end{aligned}$$

Taking the maximum with respect to $0 \leq x < g^{\max}$ and noting that $1 - \max_{0 \leq y < f^{\max}} \Pr[Z_{J^u}^u \leq 0 | Z_0^u = y] = \min_{0 \leq y < f^{\max}} \Pr[Z_{J^u}^u \geq 0 | Z_0^u = y]$, we obtain (16). Similarly, we can obtain (17). Bounds on $\mathbb{E}[P | X_0 = \frac{A}{2} - x]$ can be derived in an analogous manner, i.e., by conditioning on hitting first the bottom of the buffer or leaving the lower half of the buffer:

$$\begin{aligned} \mathbb{E} \left[P \left| X_0 = \frac{A}{2} - x \right. \right] &= \mathbb{E} [J^l | Z_0^l = -x] \\ &+ \Pr \left[Z_{J^l}^l \leq -\frac{A}{2} \left| Z_0^l = -x \right. \right] \frac{\mathbb{E} [K^l | Z_0^l = 0]}{\Pr[Z_{K^l}^l \geq \frac{A}{2} | Z_0^l = 0]}. \end{aligned}$$

Using Theorem 1, we can estimate the above quantities as follows:

$$\begin{aligned} \mathbb{E} \left[P \left| X_0 = \frac{A}{2} - x \right. \right] &= \Theta(\mathbb{E}[K^l | Z_0^l = 0]) \\ &= \Theta \left(\frac{\mathbb{E}[Z_{K^l}^l | Z_0^l = 0]}{\mathbb{E}[W_1^d(\hat{X}_k = 0)]} \right) = \Theta(\delta_d^{-1}), \end{aligned}$$

$$\Pr \left[Z_{K^l}^l \leq -\frac{A}{2} \left| Z_0^l = 0 \right. \right] = 1 - \Pr[Z_{K^l}^l \geq 0 | Z_0^l = 0] = \Theta(1),$$

$$\mathbb{E}[J^u | Z_0^u = y] = \Theta(1),$$

$$\Pr \left[Z_{J^u}^u \geq \frac{A}{2} \left| Z_0^u = y \right. \right] = \Theta \left(\exp \left(-\frac{A}{2} r_d^*(\hat{X}_k = 1) \right) \right),$$

$$\mathbb{E}[K^u | Z_0^u = 0] = \Theta(A),$$

where $\delta_d = \mathbb{E}[W_1^d(\hat{X}_k = 0)]$ is the drift in the lower half of the buffer. Inserting these estimates into (13) leads to

$$\mathbb{E}[Y] = \Theta \left(\delta_d^{-1} \exp \left(\frac{A}{2} r_d^*(\hat{X}_k = 1) \right) \right). \quad (18)$$

Lemma 3 implies

$$\begin{aligned} r_d^*(\hat{X}_k = 1) &= \frac{2\delta_d}{\text{var}(W_k^d(\hat{X}_k = 1) | \delta_d = \delta_e = 0) + O(|\delta_e|)} - O(\delta_d^2) \\ &= \frac{\ln A}{A} \frac{2\beta_1}{\text{var}(W_k^d(\hat{X}_k = 1) | \delta_d = \delta_e = 0) + O(|\delta_e|)} \\ &- O \left(\left(\frac{\ln A}{A} \right)^2 \right), \end{aligned} \quad (19)$$

where $\delta_d = \frac{\beta_1 \ln A}{A}$ and the term $O(|\delta_e|)$ accounts for the variation of the second-order statistic of the Markov chain $W_k^d(\hat{X}_k)$ around $\delta_d = \delta_e = 0$.

Combining $p_{\text{FQ}}^{\circ} = \frac{1}{\mathbb{E}[Y]}$, (18), and (19), and assuming $\beta_1/2 > \text{var}(W_k^d(\hat{X}_k) | \delta_d = \delta_e = 0) + O(|\delta_e|)$, we arrive at

$$\begin{aligned} p_{\text{FQ}}^{\circ} &= O \left(\delta_d \exp \left(-\frac{A}{2} r_d^*(\hat{X}_k = 1) \right) \right) = O \left(\frac{\ln A}{A} \right. \\ &\cdot \exp \left[\ln A \frac{-\beta_1}{\text{var}(W_k^d(\hat{X}_k = 1) | \delta_d = \delta_e = 0) + O(|\delta_e|)} \right. \\ &\left. \left. + O \left(\frac{\ln^2 A}{A} \right) \right] \right) = o \left(\frac{1}{A^2} \right). \end{aligned}$$

A similar derivation can be used to show that the probability for an empty energy buffer scales as $p_{\text{EB}}^{\circ} = o(1/B^2)$. Here, the decremental drift is chosen as $\delta_e = \frac{\beta_2 \ln B}{B}$ and we need to choose with $\beta_2/2 > \text{var}(W_k^e(\hat{E}_k) | \delta_d = \delta_e = 0) + O(|\delta_d|)$. We note that the estimates for p_{FQ}° and p_{EB}° are tight if the renewal epochs of the Markov chains \hat{X}_k and \hat{E}_k are sufficiently small compared to A and B , respectively (see [18] and references therein for further details).

We next analyze the convergence of the average distortion to the optimal value. The average distortion during normal operation as defined in (11) can be expressed via the Taylor expansion

$$\begin{aligned} D^{\text{op}} &= D_{\text{T}}(0, 0) + \frac{\partial D_{\text{T}}(0, 0)}{\partial \delta_d} \frac{\beta_1 \ln A}{A} (2q_d - 1) \\ &+ \frac{\partial D_{\text{T}}(0, 0)}{\partial \delta_e} \frac{\beta_2 \ln B}{B} (2q_e - 1) \\ &+ \frac{\partial^2 D_{\text{T}}(0, 0)}{2\partial \delta_d^2} \left(\frac{\beta_1 \ln A}{A} \right)^2 + \frac{\partial^2 D_{\text{T}}(0, 0)}{2\partial \delta_e^2} \left(\frac{\beta_2 \ln B}{B} \right)^2 \\ &+ O \left(\left(\frac{\beta_1 \ln A}{A} \right)^3 + \left(\frac{\beta_2 \ln B}{B} \right)^3 \right), \end{aligned}$$

where $q_d = \Pr(\hat{X}_k = 1)$ and $q_e = \Pr(\hat{E}_k = 1)$.

We next show that the scaling behavior of the first-order terms in this expansion is $o(1/A^2)$ and $o(1/B^2)$, respectively. Since the data buffer is finite, the difference between input and output in steady state equals zero. Mathematically,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \left[\sum_{k=0}^n W_k^d + L_n^d - U_n^d \right] = 0, \quad (20)$$

where $W_k^d = f(D_k, E_{s,k}, Q_k) - g(H_k, E_{t,k})$ is the data buffer increment process (neglecting boundary effects), L_n^d is the cumulative number of bits up to time n that have been padded with zeros due to underflow, and U_n^d is the cumulative number of bits up to time n that have been lost due to overflow. According to the strong law of large numbers for renewal-reward processes [17],

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=0}^n W_k^d = \mathbb{E}[W_1^d] = \frac{\beta_1 \ln A}{A} (2q_d - 1).$$

As to the boundary effects, there is a one-to-one correspondence between the times at which U_n^d increases and the times

$$\max_{0 \leq x < g^{\max}} \mathbb{E} \left[V \left| X_0 = \frac{A}{2} - x \right. \right] \leq \frac{\max_{0 \leq x < g^{\max}} \mathbb{E} [P | X_0 = \frac{A}{2} - x] + \max_{0 \leq y < f^{\max}} \mathbb{E} [J^u | Z_0^u = y]}{\min_{0 \leq y < f^{\max}} \Pr [Z_{J^u}^u \geq \frac{A}{2} | Z_0^u = y]} \quad (16)$$

$$\min_{0 \leq x < g^{\max}} \mathbb{E} \left[V \left| X_0 = \frac{A}{2} - x \right. \right] \geq \frac{\min_{0 \leq x < g^{\max}} \mathbb{E} [P | X_0 = \frac{A}{2} - x] + \min_{0 \leq y < f^{\max}} \mathbb{E} [J^u | Z_0^u = y]}{\max_{0 \leq y < f^{\max}} \Pr [Z_{J^u}^u \geq \frac{A}{2} | Z_0^u = y]} \quad (17)$$

that the buffer is full. Furthermore, this increase is bounded by f^{\max} , so that

$$\lim_{n \rightarrow \infty} \frac{U_n^d}{n} \leq f^{\max} p_{\text{FQ}}^{\pi} = o(1/A^2)$$

Due to the symmetry of the problem, we similarly obtain $\lim_{n \rightarrow \infty} L_n^d/n = o(1/A^2)$. Hence, (20) implies $\frac{\beta_1 \ln A}{A}(2q_d - 1) = o(1/A^2)$. The same line of arguments can be used to show $\frac{\beta_2 \ln B}{B}(2q_e - 1) = o(1/B^2)$. By combining the above intermediate results, it follows that the average distortion during normal operation reads $D^{\text{op}} = D_{\text{T}}(0, 0) + O\left(\left(\frac{\beta_1 \ln A}{A}\right)^2\right) + O\left(\left(\frac{\beta_2 \ln B}{B}\right)^2\right)$. Inserting this expression for D^{op} together with $p_{\text{FQ}}^{\pi} = o(1/A^2)$ and $p_{\text{EB}}^{\pi} = o(1/B^2)$ into the upper bound (10) of the average distortion finally confirms that \bar{D}^{π} approaches \bar{D}^* as $O((\ln A/A)^2) + O((\ln B/B)^2)$. Proposition 3 is thus proved.

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