

Compressive Nonparametric Graphical Model Selection for Time Series: A Multitask Learning Approach

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Introduction

- Consider a p -dimensional vector-valued Gaussian process

$$\mathbf{x}[n] = (x_1[n], \dots, x_p[n])^T.$$

- We model $\mathbf{x}[n]$ as stationary with summable auto-covariance function (ACF)

$$\mathbf{R}[m] := \mathbb{E}\{\mathbf{x}[m]\mathbf{x}^T[0]\}.$$

- The **spectral density matrix (SDM)** is

$$\mathbf{S}(\theta) := \sum_{m=-\infty}^{\infty} \mathbf{R}[m] \exp(-j2\pi m\theta), \text{ for } \theta \in [0, 1).$$

- We require the SDM to be sufficiently smooth, or equivalently we require small ACF moments

$$\mu^{(h)} := \sum_{m=-\infty}^{\infty} h[m] \|\mathbf{R}[m]\|_{\infty}. \quad (1)$$

- Consider the **conditional independence graph (CIG)** $\mathcal{G} = ([p], E)$ of $\mathbf{x}[n]$.

- For a Gaussian process, $(k, l) \notin E \iff [\mathbf{S}^{-1}(\theta)]_{k,l} \equiv 0$.

- Our interest is in sparse CIGs, containing few edges.

- Goal:** Estimate sparse CIG from a finite-length observation $\mathbf{x}[1], \dots, \mathbf{x}[N]$, where typically $N \ll p$.

- We propose a nonparametric compressive selection scheme for the CIG of a stationary vector process.

Neighborhood Regression

- First, consider the special case of an i.i.d. sampling process, i.e., $\mathbf{R}[m] = \mathbf{C}\delta[m]$.

- This corresponds to model selection for a Gaussian Markov random field (GMRF) with covariance matrix \mathbf{C} .

- Here, the SDM is flat, i.e., $\mathbf{S}(\theta) = \mathbf{C}$ for all $\theta \in [0, 1)$.

- Determine neighborhood $\mathcal{N}(r) := \{l : (r, l) \in E\}$ by regressing $x_r[n]$ on the remaining components.

- For the i.i.d. case, this regression becomes

$$x_r[n] = \sum_{l \in [p] \setminus \{r\}} \beta_l x_l[n] + w[n],$$

$$\text{with } \beta_l = -[\mathbf{C}^{-1}]_{r,l} / [\mathbf{C}^{-1}]_{r,r}.$$

- Since $\mathcal{N}(r)$ coincides with $\text{supp}(\boldsymbol{\beta})$, with $\boldsymbol{\beta} := \text{vec}\{\beta_l\}_{l \in [p] \setminus \{r\}}$ determining the neighborhood is reduced to **sparse support recovery!**

- LASSO based selection scheme proposed by [Meinshausen & Bühlmann, 2006].

Multitask Learning Formulation

- For a general stationary process, we perform **neighborhood regression in the frequency domain**.

- Let $\hat{\mathbf{S}}(\theta_f)$ denote an estimate of the SDM $\mathbf{S}(\theta_f)$ for $\theta_f := (f-1)/F$, $f \in [F]$.

- For each frequency θ_f , $f \in [F]$, we define

$$\mathbf{y}^{(f)} = \mathbf{X}^{(f)} \boldsymbol{\beta}^{(f)} + \mathbf{w}^{(f)}, \quad (2)$$

with

$$(\mathbf{y}^{(f)} \ \mathbf{X}^{(f)}) := \sqrt{\mathbf{P}_{1 \leftrightarrow r} \hat{\mathbf{S}}(\theta_f) \mathbf{P}_{1 \leftrightarrow r}}, \quad (3)$$

and parameter vectors

$$\boldsymbol{\beta}^{(f)} := [\mathbf{S}(\theta_f)]_{[p] \setminus \{r\}, [p] \setminus \{r\}}^{-1} [\mathbf{S}(\theta_f)]_{[p] \setminus \{r\}, r}. \quad (4)$$

- Based on (4) it can be shown that

$$\mathcal{N}(r) = \text{gsupp}(\boldsymbol{\beta}) := \bigcup_{f \in [F]} \text{supp}(\boldsymbol{\beta}^{(f)}),$$

with stacked parameter vector $\boldsymbol{\beta} := ((\boldsymbol{\beta}^{(1)})^T, \dots, (\boldsymbol{\beta}^{(F)})^T)^T$.

- If the CIG is sparse, i.e., $|\mathcal{N}(r)| \ll p$, it follows that $\boldsymbol{\beta}$ is a block-sparse vector.

- Recovering a block-sparse vector $\boldsymbol{\beta}$ from the measurements (2) is recognized as a **multitask learning problem**.

- A popular approach to this is **multitask LASSO (mLASSO)**.

Novel Selection Scheme

- Let $w[m]$ denote a window function with non-negative discrete time Fourier transform.

- We propose the following selection scheme:

- First, for each θ_f , we compute a multivariate **Blackman-Tukey SDM estimate**

$$\hat{\mathbf{S}}(\theta) := \sum_{m=-N+1}^{N-1} w[m] \hat{\mathbf{R}}[m] \exp(-j2\pi m\theta).$$

Here, $\hat{\mathbf{R}}[m] := (1/N) \sum_{n=1}^{N-m} \mathbf{x}[n+m] \mathbf{x}^T[n]$ for $m \in \{0, \dots, N-1\}$ and, by symmetry of the ACF, $\hat{\mathbf{R}}[m] := \hat{\mathbf{R}}^T[-m]$ for $m \in \{-N+1, \dots, -1\}$.

- Second, based on $\hat{\mathbf{S}}(\theta_f)$, $f \in [F]$, we construct $\mathbf{y}^{(f)}$ and $\mathbf{X}^{(f)}$ according to (3) and compute mLASSO estimate

$$\hat{\boldsymbol{\beta}} = \underset{\boldsymbol{\beta}}{\text{argmin}} \left\{ \frac{1}{F} \sum_{f \in [F]} \|\mathbf{y}^{(f)} - \mathbf{X}^{(f)} \boldsymbol{\beta}^{(f)}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_{2,1} \right\},$$

where $\|\boldsymbol{\beta}\|_{2,1} := \sum_{r \in [q]} \|\boldsymbol{\beta}_r\|_2$ with $\boldsymbol{\beta}_r := ([\boldsymbol{\beta}^{(1)}]_r, \dots, [\boldsymbol{\beta}^{(F)}]_r)^T \in \mathbb{C}^F$.

- The neighborhood $\mathcal{N}(r)$ is finally estimated by

$$\hat{\mathcal{N}}(r) := \{l : \|\hat{\boldsymbol{\beta}}_l\|_2 > \tau\}.$$

- Scheme is modular: Different combinations of SDM estimators and sparse support recovery schemes possible.

Performance Guarantees

- Denote by $\mathcal{M}(A, B, s_{\max}, \rho_{\min}, \mu^{(h_1)}, \phi_{\min})$ the class of stationary Gaussian processes that satisfy the following conditions

- **Uniform boundedness of SDM eigenvalues:**

$$0 < A \leq \lambda_{\min}(\mathbf{S}(\theta)) \leq \lambda_{\max}(\mathbf{S}(\theta)) \leq B < \infty.$$

This technical assumption ensures certain Markov properties of the CIG.

- **Maximum node degree s_{\max} :** We consider sparse CIGs, whose maximum node degree is bounded as

$$|\mathcal{N}(r)| \leq s_{\max} \ll p.$$

- **Minimum partial coherence $\rho_{\min} > 0$:** This parameter quantifies the minimum partial correlation between the spectral components of the process. In particular, we require

$$\sum_{f \in [F]} |[\mathbf{S}^{-1}(\theta_f)]_{r,r'} / [\mathbf{S}^{-1}(\theta_f)]_{r,r}|^2 \geq \rho_{\min}^2$$

for all $r \in [p]$, $r' \in \mathcal{N}(r)$.

- **ACF moment $\mu^{(h_1)}$:** We quantify the smoothness of the processes in \mathcal{M} using the ACF moment (1) with weight function $h_1[m] := |1 - w[m](1 - |m|/N)|$.

- **Minimum multitask compatibility constant $\phi_{\min} > 0$:** For every process in \mathcal{M} , we require

$$s_{\max} \sum_{f \in [F]} (\boldsymbol{\beta}^{(f)})^H \mathbf{S}(\theta_f) \boldsymbol{\beta}^{(f)} \geq \phi_{\min}^2 \|\boldsymbol{\beta}_{\mathcal{N}(r)}\|_{2,1}^2$$

for all $r \in [p]$ and all vectors $\boldsymbol{\beta} \in \mathbb{C}^{qF}$ such that

$$\|\boldsymbol{\beta}_S\|_{2,1} > 0 \text{ and } \|\boldsymbol{\beta}_{S^c}\|_{2,1} \leq 3\|\boldsymbol{\beta}_S\|_{2,1}.$$

- We choose mLASSO parameter $\lambda = \phi_{\min}^2 \rho_{\min} / (18s_{\max} F)$ and threshold $\tau = \rho_{\min} / 2$.

- Combining a deterministic analysis of mLASSO with a large-deviation characterization of the SDM estimator, we derived the following result:

Theorem 1 Consider a p -dimensional stationary Gaussian time series $\mathbf{x}[n]$ belonging to $\mathcal{M}(A, B, s_{\max}, \rho_{\min}, \mu^{(h_1)}, \phi_{\min})$. Then, if for some $\delta > 0$, the rescaled sample size $\eta := N / (\log(p) s_{\max}^3)$ and the correlation moment $\mu^{(h_1)}$ satisfy

$$\eta > 10^3 \log\left(\frac{4F}{\delta}\right) \|w\|_1^2 B^2 / \kappa^2 \text{ and } \mu^{(h_1)} \leq \frac{\kappa}{2s_{\max}^{3/2}},$$

with $\kappa := (\phi_{\min}^2 / 174) \rho_{\min} \sqrt{AF/B}$, the probability of correctly selecting the edge set E is at least $1 - \delta$, i.e.,

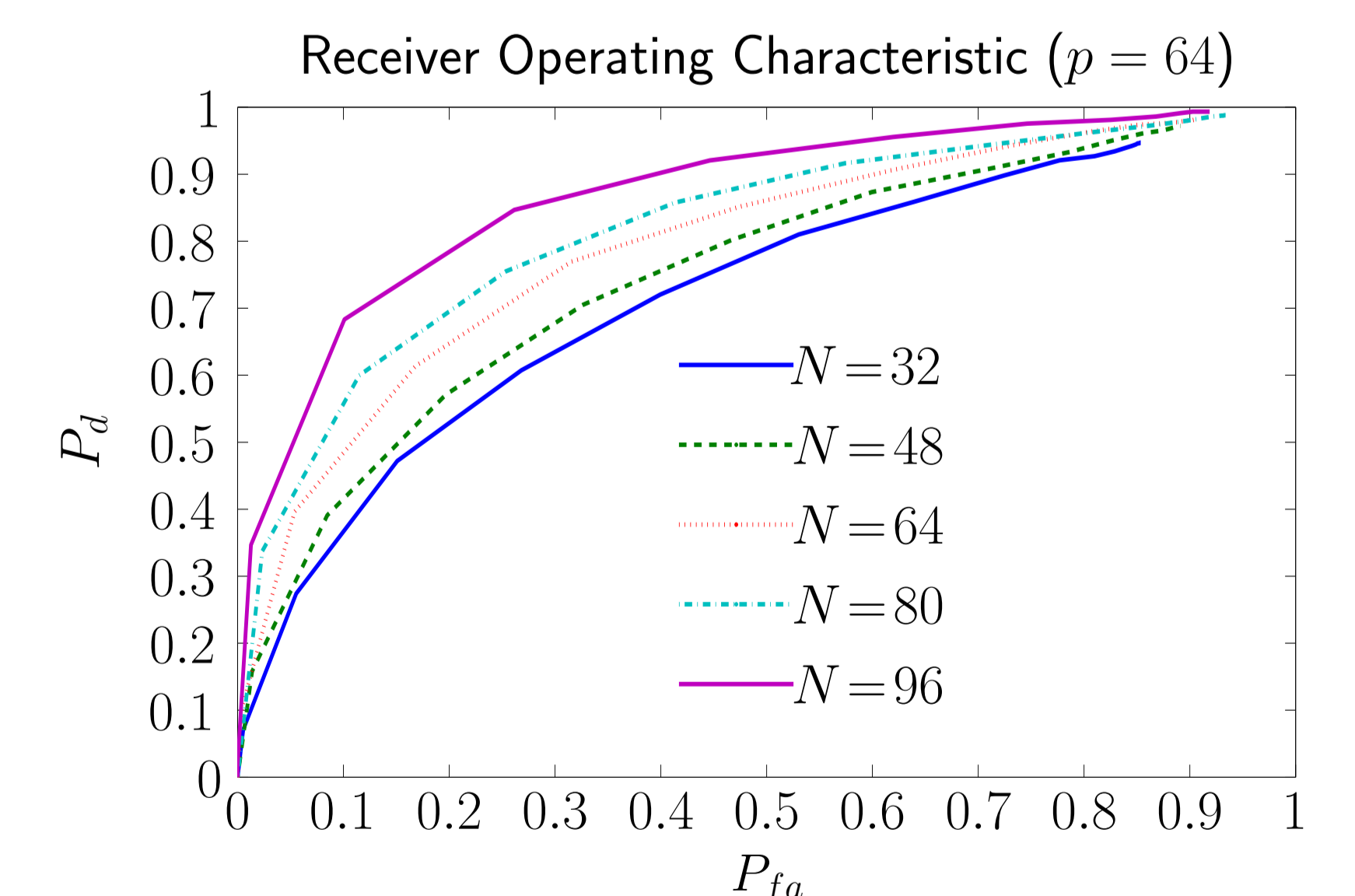
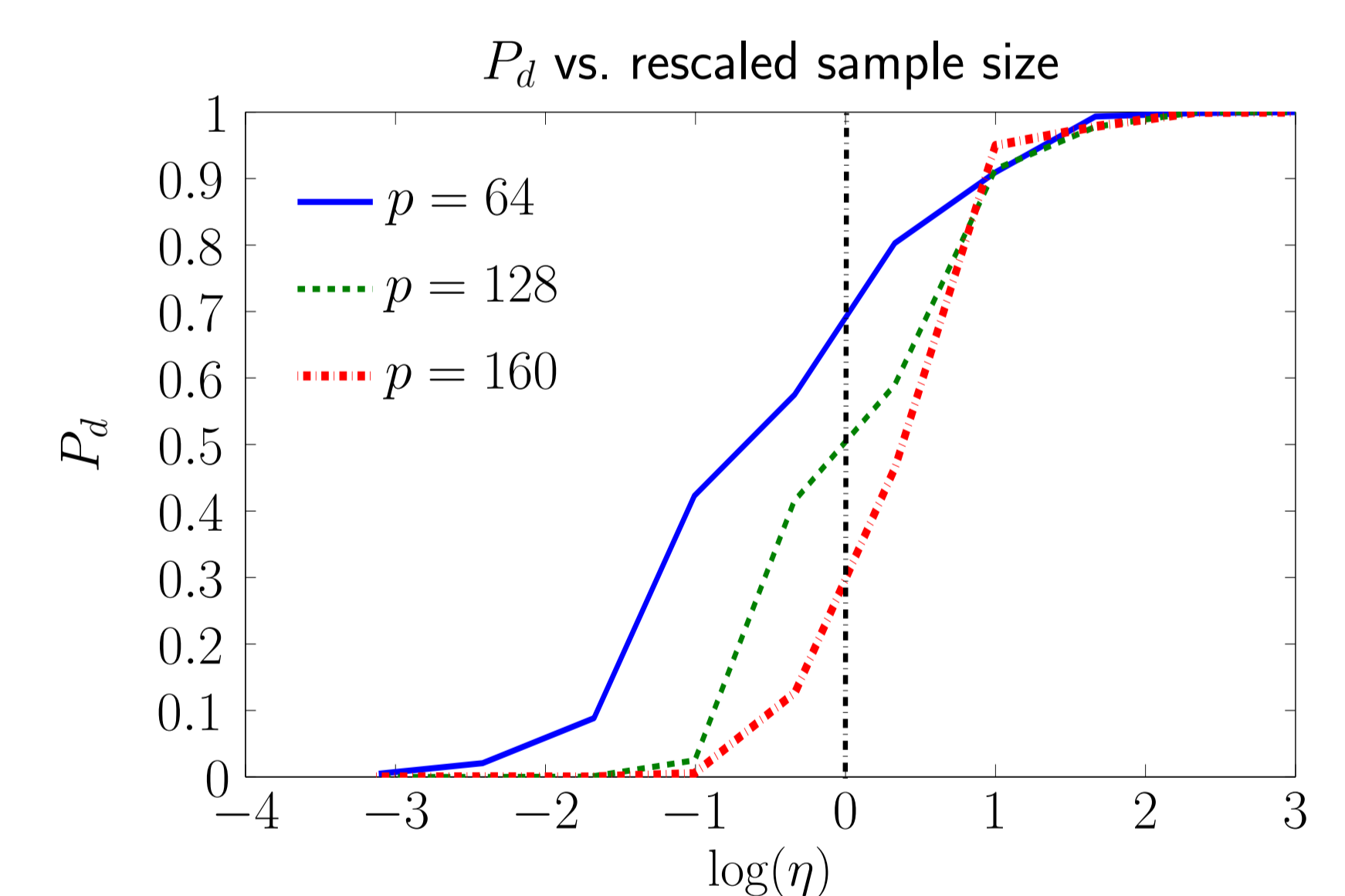
$$\mathbb{P}\left\{ \bigcap_{r \in [p]} \{\hat{\mathcal{N}}(r) = \mathcal{N}(r)\} \right\} \geq 1 - \delta.$$

Numerical Results

- We applied our method to a synthetic process obtained by filtering a p -dimensional white Gaussian process with a FIR filter of length 2.

- The filtered process has a sparse CIG with $s_{\max} = 3$.

- We computed empirical false alarm (P_{fa}) and detection ratios (P_d) based on 10 simulation runs.



- Our method yields reasonable performance even if $N = 32$ only for a 64-dimensional process.

- The rescaled sample size η seems to be a good performance indicator.