

Sigma Point Belief Propagation

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- 1 Sigma Point Basics
- 2 Bayesian Estimation with Sigma Points
- 3 Belief Propagation
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- Consider a **random vector** $\mathbf{x} \in \mathbb{R}^J$ whose mean $\boldsymbol{\mu}_x$ and covariance matrix \mathbf{C}_x are known, and a **transformed random vector** $\mathbf{y} = H(\mathbf{x})$.

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- **Sigma points (SPs)** [Julier et al., 1997] are used to calculate approximations of the mean $\boldsymbol{\mu}_y$, covariance matrix \mathbf{C}_y , and cross-covariance matrix $\mathbf{C}_{x,y}$.
- The resulting approximations are at least as good as those obtained by linearizing $H(\cdot)$.
- Fundamental difference to Monte Carlo type methods: The SPs are not random samples but are calculated by means of a **deterministic** algorithm.

Calculation of sigma points

- For a J -dimensional random vector \mathbf{x} , SPs and corresponding weights $\{(\mathbf{x}^{(j)}, w^{(j)})\}_{j=0}^{2J}$ are calculated as [Julier et al., 1997]

$$\mathbf{x}^{(j)} = \begin{cases} \boldsymbol{\mu}_{\mathbf{x}}, & j = 0 \\ \boldsymbol{\mu}_{\mathbf{x}} + \sqrt{J + \kappa} (\mathbf{C}_{\mathbf{x}}^{1/2})_j, & j = 1, \dots, J \\ \boldsymbol{\mu}_{\mathbf{x}} - \sqrt{J + \kappa} (\mathbf{C}_{\mathbf{x}}^{1/2})_j, & j = J + 1, \dots, 2J \end{cases}$$
$$w^{(j)} = \begin{cases} \frac{\kappa}{J + \kappa}, & j = 0 \\ \frac{1}{2(J + \kappa)}, & j = 1, \dots, 2J. \end{cases}$$

Here, $(\mathbf{C}_{\mathbf{x}}^{1/2})_j$ is the j th row or column of the matrix square root of $\mathbf{C}_{\mathbf{x}}$ and κ is a parameter controlling the spread of the SPs around the mean $\boldsymbol{\mu}_{\mathbf{x}}$.

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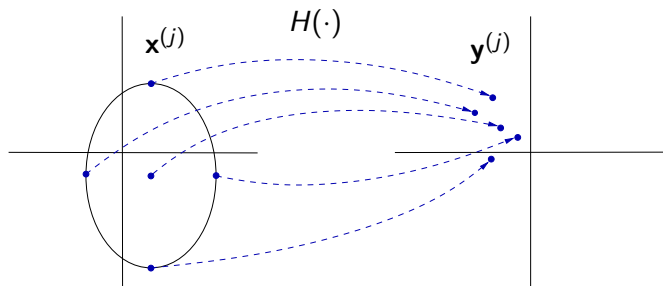
Here, $(\mathbf{C}_{\mathbf{x}}^{1/2})_j$ is the j th row or column of the matrix square root of $\mathbf{C}_{\mathbf{x}}$ and κ is a parameter controlling the spread of the SPs around the mean $\boldsymbol{\mu}_{\mathbf{x}}$.

- SPs have the property that the **weighted sample mean** $\tilde{\boldsymbol{\mu}}_{\mathbf{x}} = \sum_{j=0}^{2J} w^{(j)} \mathbf{x}^{(j)}$ and **weighted sample covariance matrix** $\tilde{\mathbf{C}}_{\mathbf{x}} = \sum_{j=0}^{2J} w^{(j)} (\mathbf{x}^{(j)} - \tilde{\boldsymbol{\mu}}_{\mathbf{x}})(\mathbf{x}^{(j)} - \tilde{\boldsymbol{\mu}}_{\mathbf{x}})^{\top}$ are exactly equal to $\boldsymbol{\mu}_{\mathbf{x}}$ and $\mathbf{C}_{\mathbf{x}}$, respectively.

Unscented transformation

- SPs $\{\mathbf{y}^{(j)}\}_{j=0}^{2J}$ of \mathbf{y} can be obtained by propagating each SP $\mathbf{x}^{(j)}$ through $H(\cdot)$ [Julier et al., 1997]:

$$\mathbf{y}^{(j)} = H(\mathbf{x}^{(j)}), \quad j = 0, \dots, 2J.$$



Calculation of mean and covariance

- From $\{(\mathbf{x}^{(j)}, \mathbf{y}^{(j)}, w^{(j)})\}_{j=0}^{2J}$, one can calculate approximations of $\mu_{\mathbf{y}}$, $\mathbf{C}_{\mathbf{y}}$ and $\mathbf{C}_{\mathbf{xy}}$ as

$$\tilde{\mu}_{\mathbf{y}} = \sum_{j=0}^{2J} w^{(j)} \mathbf{y}^{(j)},$$

$$\tilde{\mathbf{C}}_{\mathbf{y}} = \sum_{j=0}^{2J} w^{(j)} (\mathbf{y}^{(j)} - \mu_{\mathbf{y}})(\mathbf{y}^{(j)} - \mu_{\mathbf{y}})^{\top},$$

$$\tilde{\mathbf{C}}_{\mathbf{xy}} = \sum_{j=0}^{2J} w^{(j)} (\mathbf{y}^{(j)} - \mu_{\mathbf{y}})(\mathbf{x}^{(j)} - \mu_{\mathbf{x}})^{\top}.$$

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- SPs can be used for **Bayesian estimation** of a random vector \mathbf{x} from an observed vector

$$\mathbf{z} = \mathbf{y} + \mathbf{n}, \quad \text{with } \mathbf{y} = H(\mathbf{x}).$$

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- Bayesian estimation relies on the **posterior probability density function** (pdf)

$$f(\mathbf{x}|\mathbf{z}) \propto f(\mathbf{z}|\mathbf{x})f(\mathbf{x}),$$

where $f(\mathbf{z}|\mathbf{x})$ is the likelihood function and $f(\mathbf{x})$ is the prior pdf.

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- Closed-form calculation of $f(\mathbf{x}|\mathbf{z})$ is usually infeasible!

- A feasible special case is when $H(\cdot)$ is linear, i.e., $H(\mathbf{x}) = \mathbf{H}\mathbf{x}$ and \mathbf{x} and \mathbf{n} are Gaussian. Then $f(\mathbf{x}|\mathbf{z})$ is also Gaussian, and $\boldsymbol{\mu}_{\mathbf{x}|\mathbf{z}}$ and $\mathbf{C}_{\mathbf{x}|\mathbf{z}}$ can be calculated as

$$\boldsymbol{\mu}_{\mathbf{x}|\mathbf{z}} = \boldsymbol{\mu}_{\mathbf{x}} + \mathbf{K}(\mathbf{z} - \boldsymbol{\mu}_{\mathbf{y}}), \quad \mathbf{C}_{\mathbf{x}|\mathbf{z}} = \mathbf{C}_{\mathbf{x}} - \mathbf{K}(\mathbf{C}_{\mathbf{y}} + \mathbf{C}_{\mathbf{n}})\mathbf{K}^T,$$

with

$$\begin{aligned}\mathbf{K} &= \mathbf{C}_{\mathbf{xy}}(\mathbf{C}_{\mathbf{y}} + \mathbf{C}_{\mathbf{n}})^{-1}, \\ \boldsymbol{\mu}_{\mathbf{y}} &= \mathbf{H}\boldsymbol{\mu}_{\mathbf{x}}, \\ \mathbf{C}_{\mathbf{y}} &= \mathbf{H}\mathbf{C}_{\mathbf{x}}\mathbf{H}^T, \\ \mathbf{C}_{\mathbf{xy}} &= \mathbf{C}_{\mathbf{x}}\mathbf{H}^T.\end{aligned}$$

Bayesian estimation with sigma points

- In the **nonlinear case**, $\mu_{x|z}$ and $\mathbf{C}_{x|z}$ can be approximated by means of SPs [Julier et al., 1997].
- This is done by using the closed-form expressions of the linear-Gaussian case, in which μ_y , \mathbf{C}_y , and \mathbf{C}_{xy} are replaced by the SP approximations $\tilde{\mu}_y$, $\tilde{\mathbf{C}}_y$, and $\tilde{\mathbf{C}}_{xy}$.

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- SP approximations $\tilde{\mu}_{x|z}$ and $\tilde{\mathbf{C}}_{x|z}$ are thus obtained as

$$\tilde{\mu}_{x|z} = \mu_x + \tilde{\mathbf{K}}(z - \tilde{\mu}_y), \quad \tilde{\mathbf{C}}_{x|z} = \mathbf{C}_x - \tilde{\mathbf{K}}(\tilde{\mathbf{C}}_y + \mathbf{C}_n)\tilde{\mathbf{K}}^T,$$

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- **Sequential implementation** of this algorithm \Rightarrow **SP filter**

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Here, $\mathbf{z}_{k,l} = H(\mathbf{x}_k, \mathbf{x}_l) + \mathbf{v}_{k,l}$ are noisy “pairwise” observations.

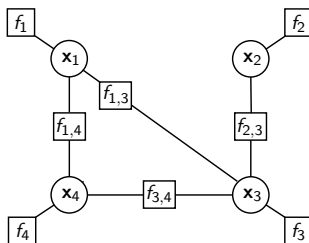
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- Representation by factor graph:



○ variable node

□ factor node

$k \in \{1, 2, 3, 4\}$

$\mathcal{E} = \{(1, 3), (1, 4), (2, 3), (3, 4)\}$

$f_k \triangleq f(\mathbf{x}_k)$

$f_{k,l} \triangleq f(\mathbf{z}_{k,l} | \mathbf{x}_k, \mathbf{x}_l)$

- To obtain an estimate of \mathbf{x}_k , the “marginal” posterior pdf $f(\mathbf{x}_k|\mathbf{z}) = \int f(\mathbf{x}|\mathbf{z}) d\mathbf{x}^{\sim k}$ is needed.
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- An approximate marginal posterior (“belief”) $b(\mathbf{x}_k) \approx f(\mathbf{x}_k|\mathbf{z})$ is obtained by running belief propagation (BP) message passing on the factor graph representing the joint posterior $f(\mathbf{x}|\mathbf{z})$.
- In the special case where the factor graph is a tree (no loops/cycles), BP is exact, i.e., $b(\mathbf{x}_k) = f(\mathbf{x}_k|\mathbf{z})$.

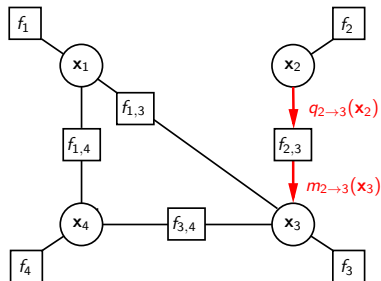
Belief propagation rules

- Let \mathcal{N}_k denote the “neighbor” set of variable node $k \in \{1, \dots, K\}$, which comprises all variable nodes $l \in \{1, \dots, K\} \setminus \{k\}$ such that $(k, l) \in \mathcal{E}$.

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- Belief of variable \mathbf{x}_k :

$$b(\mathbf{x}_k) \propto f(\mathbf{x}_k) \prod_{l \in \mathcal{N}_k} m_{l \rightarrow k}(\mathbf{x}_k).$$



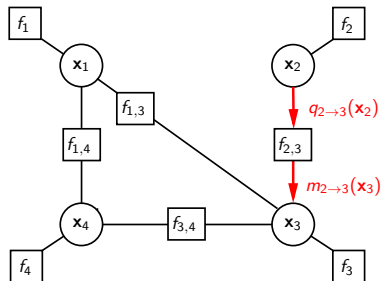
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- **Message** from variable node \mathbf{x}_k to function node $f(\mathbf{z}_{k,l} | \mathbf{x}_k, \mathbf{x}_l)$:

$$q_{k \rightarrow l}(\mathbf{x}_k) = f(\mathbf{x}_k) \prod_{l' \in \mathcal{N}_k \setminus \{l\}} m_{l' \rightarrow k}(\mathbf{x}_k).$$



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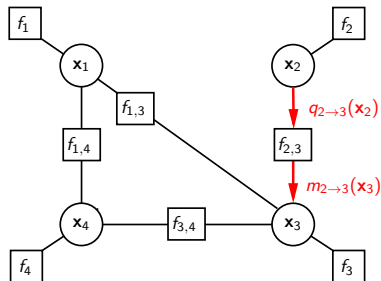
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$$m_{l \rightarrow k}(\mathbf{x}_k) = \int f(\mathbf{z}_{k,l} | \mathbf{x}_k, \mathbf{x}_l) q_{l \rightarrow k}(\mathbf{x}_l) d\mathbf{x}_l.$$



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- In this work, we develop **sigma point belief propagation (SPBP)** by extending the SP filter to Bayesian inference corresponding to a general factorization structure.

Factorization Structure ← System Properties ↓	Sequential	General
Linear, Gaussian	Kalman filter [Kalman, 1960]	Gaussian BP [Weiss et al., 2001]
Nonlinear, non-Gaussian	Particle filter [Gordon et al., 1993]	Nonparametric BP [Sudderth et al., 2003]
Nonlinear, Gaussian	SP filter [Julier et al., 1997]	SPBP

Reformulation of BP rules

- Recall BP rules:

$$b(\mathbf{x}_k) \propto f(\mathbf{x}_k) \prod_{l \in \mathcal{N}_k} m_{l \rightarrow k}(\mathbf{x}_k)$$

$$q_{k \rightarrow l}(\mathbf{x}_k) = f(\mathbf{x}_k) \prod_{l' \in \mathcal{N}_k \setminus \{l\}} m_{l' \rightarrow k}(\mathbf{x}_k)$$

$$m_{l \rightarrow k}(\mathbf{x}_k) = \int f(\mathbf{z}_{k,l} | \mathbf{x}_k, \mathbf{x}_l) q_{l \rightarrow k}(\mathbf{x}_l) d\mathbf{x}_l.$$

- Recall BP rules:

$$b(\mathbf{x}_k) \propto f(\mathbf{x}_k) \prod_{I \in \mathcal{N}_k} m_{I \rightarrow k}(\mathbf{x}_k)$$

$$q_{k \rightarrow I}(\mathbf{x}_k) = f(\mathbf{x}_k) \prod_{I' \in \mathcal{N}_k \setminus \{I\}} m_{I' \rightarrow k}(\mathbf{x}_k)$$

$$m_{I \rightarrow k}(\mathbf{x}_k) = \int f(\mathbf{z}_{k,I} | \mathbf{x}_k, \mathbf{x}_I) q_{I \rightarrow k}(\mathbf{x}_I) d\mathbf{x}_I.$$

- Equivalently,

$$b(\mathbf{x}_k) \propto \int f(\mathbf{x}_k) \prod_{I \in \mathcal{N}_k} [f(\mathbf{z}_{k,I} | \mathbf{x}_k, \mathbf{x}_I) q_{I \rightarrow k}(\mathbf{x}_I) d\mathbf{x}_I]$$

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Reformulation of BP rules

- Let $\mathcal{N}_k = \{l_1, l_2, \dots, l_{|\mathcal{N}_k|}\}$, and consider the “composite” vectors

$$\bar{\mathbf{x}}_k \triangleq (\mathbf{x}_k^\top \mathbf{x}_{l_1}^\top \mathbf{x}_{l_2}^\top \cdots \mathbf{x}_{l_{|\mathcal{N}_k|}}^\top)^\top \quad (\mathbf{x}_k \text{ and its neighbor states}),$$

$$\bar{\mathbf{z}}_k \triangleq (\mathbf{z}_{k,l_1}^\top \mathbf{z}_{k,l_2}^\top \cdots \mathbf{z}_{k,l_{|\mathcal{N}_k|}}^\top)^\top \quad (\text{all observations involving } \mathbf{x}_k).$$

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- We can now write

$$\begin{aligned} b(\mathbf{x}_k) &\propto \int f(\mathbf{x}_k) \prod_{l \in \mathcal{N}_k} [f(\mathbf{z}_{k,l} | \mathbf{x}_k, \mathbf{x}_l) q_{l \rightarrow k}(\mathbf{x}_l) d\mathbf{x}_l] \\ &\propto \int f(\bar{\mathbf{z}}_k | \bar{\mathbf{x}}_k) f(\bar{\mathbf{x}}_k) d\bar{\mathbf{x}}_k^{\sim k}, \end{aligned}$$

with

$$f(\bar{\mathbf{x}}_k) \propto f(\mathbf{x}_k) \prod_{l \in \mathcal{N}_k} q_{l \rightarrow k}(\mathbf{x}_l) \quad (\text{composite prior}),$$

$$f(\bar{\mathbf{z}}_k | \bar{\mathbf{x}}_k) = \prod_{l \in \mathcal{N}_k} f(\mathbf{z}_{k,l} | \mathbf{x}_k, \mathbf{x}_l) \quad (\text{composite likelihood}).$$

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- Note that $f(\bar{\mathbf{z}}_k | \bar{\mathbf{x}}_k)$ corresponds to the composite observation model

$$\bar{\mathbf{z}}_k = \bar{\mathbf{y}}_k + \bar{\mathbf{v}}_k, \quad \text{with } \bar{\mathbf{y}}_k = H(\bar{\mathbf{x}}_k),$$

where $H(\bar{\mathbf{x}}_k) \triangleq ((G(\mathbf{x}_k, \mathbf{x}_{l_1}))^\top \cdots (G(\mathbf{x}_k, \mathbf{x}_{l_{|\mathcal{N}_k|}}))^\top)^\top$ and $\bar{\mathbf{v}}_k \triangleq (\mathbf{v}_{k,l_1}^\top \cdots \mathbf{v}_{k,l_{|\mathcal{N}_k|}}^\top)^\top$.

- We can finally express $b(\mathbf{x}_k)$ as the result of a **marginalization**:

$$b(\mathbf{x}_k) = \int b(\bar{\mathbf{x}}_k) d\bar{\mathbf{x}}_k^{\sim k},$$

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- This expression of $b(\bar{\mathbf{x}}_k)$ has the same form as the expression $f(\mathbf{x}|\mathbf{z}) \propto f(\mathbf{z}|\mathbf{x})f(\mathbf{x})$ that occurred in Bayesian estimation \Rightarrow SPs can again be used for calculating an approximate mean and covariance of $b(\bar{\mathbf{x}}_k)$ and, in turn, of $b(\mathbf{x}_k)$.

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- More specifically, $\tilde{\boldsymbol{\mu}}_{b(\bar{\mathbf{x}}_k)}$ and $\tilde{\mathbf{C}}_{b(\bar{\mathbf{x}}_k)}$ can be obtained by using the closed-form expressions of the linear-Gaussian case, in which $\boldsymbol{\mu}_{\bar{\mathbf{y}}_k}$, $\mathbf{C}_{\bar{\mathbf{y}}_k}$, and $\mathbf{C}_{\bar{\mathbf{x}}_k \bar{\mathbf{y}}_k}$ are replaced by the SP approximations $\tilde{\boldsymbol{\mu}}_{\bar{\mathbf{y}}_k}$, $\tilde{\mathbf{C}}_{\bar{\mathbf{y}}_k}$, and $\tilde{\mathbf{C}}_{\bar{\mathbf{x}}_k \bar{\mathbf{y}}_k}$.

SPBP algorithm

An SP-based approximate calculation of the mean $\mu_{b(\mathbf{x}_k)}$ and covariance matrix $\mathbf{C}_{b(\mathbf{x}_k)}$ of $b(\mathbf{x}_k)$ can be obtained by performing the following two steps:

- 1 Use the “Bayesian estimation” SP scheme to calculate $\tilde{\mu}_{b(\bar{\mathbf{x}}_k)} \approx \mu_{b(\bar{\mathbf{x}}_k)}$ and $\tilde{\mathbf{C}}_{b(\bar{\mathbf{x}}_k)} \approx \mathbf{C}_{b(\bar{\mathbf{x}}_k)}$ representing $b(\bar{\mathbf{x}}_k)$ from $\mu_{\bar{\mathbf{x}}_k}$ and $\mathbf{C}_{\bar{\mathbf{x}}_k}$ representing $f(\bar{\mathbf{x}}_k)$.

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- 2 Obtain $\tilde{\mu}_{b(\mathbf{x}_k)}$ and $\tilde{\mathbf{C}}_{b(\mathbf{x}_k)}$ by extracting from $\tilde{\mu}_{b(\bar{\mathbf{x}}_k)}$ and $\tilde{\mathbf{C}}_{b(\bar{\mathbf{x}}_k)}$ the elements related to \mathbf{x}_k . This corresponds to the marginalization

$$b(\mathbf{x}_k) = \int b(\bar{\mathbf{x}}_k) d\bar{\mathbf{x}}_k^{\sim k}.$$

- 1 Sigma Point Basics
- 2 Bayesian Estimation with Sigma Points
- 3 Belief Propagation
- 4 Sigma Point Belief Propagation
- 5 Simulation Results**

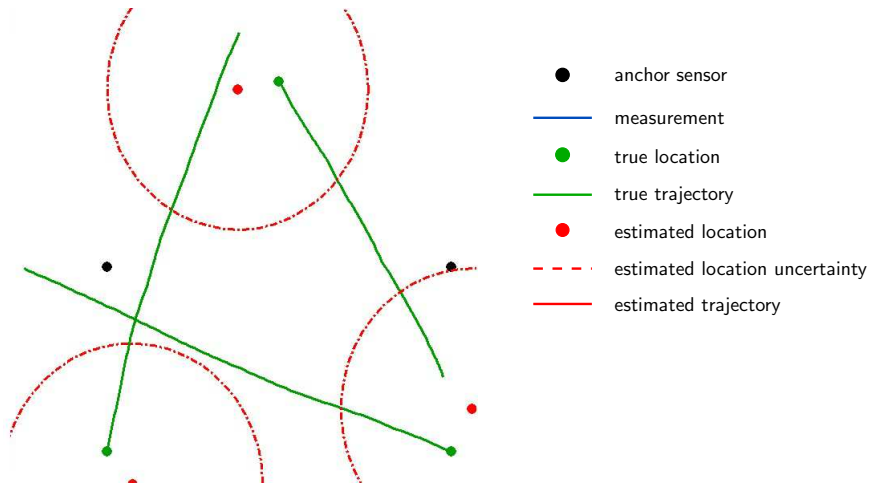
Cooperative self-localization

- We simulated a decentralized, cooperative, dynamic localization scenario [Wymeersch et al., 2009] using a network of three mobile sensors and two anchor sensors.
- The state $\mathbf{x}_{k,i}$ of mobile sensor $k \in \{1, 2, 3\}$ at time $i \in \{0, 1, \dots, 50\}$ consists of the sensor's location and velocity and corresponds to a variable node in the factor graph.

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- Each mobile sensor communicates with all other sensors, performs distance measurements relative to all other sensors, and estimates its own state.
- The communication graph coincides with the factor graph; this facilitates a distributed processing.

Simulation scenario



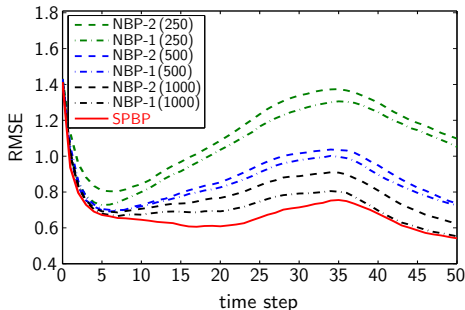
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Root-mean-square error (RMSE) of location and velocity versus time:



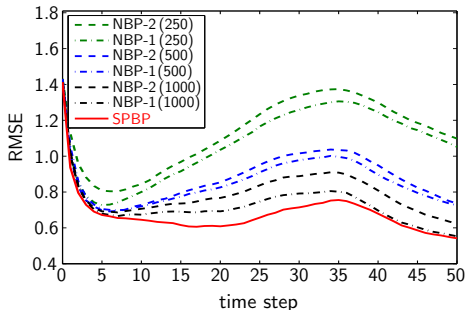
Communication and computation requirements:

	Communication between sensors [real values]	Runtime on Xeon X5650 [seconds]
SPBP	500	0.61
NBP-1 (250)	50000	1.53
NBP-2 (250)	50000	2.01
NBP-1 (500)	100000	5.16
NBP-2 (500)	100000	7.27
NBP-1 (1000)	200000	19.57
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- The proposed SPBP algorithm outperforms NBP-1 and NBP-2 although it requires significantly less communications and computations.

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- SPBP is well suited to certain decentralized inference problems in wireless sensor networks because of its **low communication requirements**.
- In a cooperative self-localization scenario, SPBP can **outperform nonparametric BP** although it requires significantly less communications and computations.

Thank you!

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