

Information Bottleneck on General Alphabets

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Abstract—We prove rigorously a source coding theorem that can probably be considered folklore, a generalization to arbitrary alphabets of a problem motivated by the Information Bottleneck method. For general random variables (Y, X) , we show essentially that for some $n \in \mathbb{N}$, a function f with rate limit $\log|f| \leq nR$ and $I(Y^n; f(X^n)) \geq nS$ exists if and only if there is a random variable U such that the Markov chain $Y \dashv\vdash X \dashv\vdash U$ holds, $I(U; X) \leq R$ and $I(U; Y) \geq S$. The proof relies on the well established discrete case and showcases a technique for lifting discrete coding theorems to arbitrary alphabets.

I. INTRODUCTION

Since its inception [1], the *Information Bottleneck* (IB) method became a widely applied tool, especially in the context of machine learning problems. It has been successfully applied to various problems in machine learning [2], computer vision [3], and communications [4], [5], [6]. Furthermore, it is a valuable tool for channel output compression in a communication system [7], [8].

In the underlying information-theoretic problem, we define a pair $(S, R) \in \mathbb{R}^2$ to be *achievable* for the two arbitrary random sources (Y, X) , if there exists a function f with rate limited range $\frac{1}{n} \log|f| \leq R$ and $I(\mathbf{Y}; f(\mathbf{X})) \geq nS$, where (\mathbf{Y}, \mathbf{X}) are n independent and identically distributed (i.i.d.) copies of (Y, X) .

While this Shannon-theoretic problem and variants thereof were also considered (e. g., [9], [10]), a large part of the literature is aimed at studying the IB function

$$S_{\text{IB}}(R) = \sup_{\substack{U : I(U; X) \leq R \\ Y \dashv\vdash X \dashv\vdash U}} I(U; Y) \quad (1)$$

in different contexts. In particular, several works (e. g., [1], [2], [11], [12], [13]) intend to compute a probability distribution that achieves the supremum in (1). The resulting distribution is then used as a building block in numerical algorithms, e. g., for document clustering [2] or dimensionality reduction [11].

In the discrete case, $S_{\text{IB}}(R)$ is equal to the maximum of all S such that (S, R) is in the *achievable region* (closure of the set of all achievable pairs). This statement has been re-proven many times in different contexts [14], [10], [15], [16]. In this note, we prove a theorem, which can probably be considered folklore, extending this result from discrete to arbitrary random variables. Formally speaking, using the definitions in [17], we prove that a pair (S, R) is in the achievable region of an arbitrary source (Y, X) if and only if, for every $\varepsilon > 0$, there exists a random variable U with $Y \dashv\vdash X \dashv\vdash U$, $I(X; U) \leq R + \varepsilon$, and $I(Y; U) \geq S - \varepsilon$. This provides a single-letter solution to the information-theoretic problem behind the information bottleneck method for arbitrary random sources and in particular it shows, that the information bottleneck for Gaussian random variables [11] is indeed the solution to a Shannon-theoretic problem.

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The proof relies on the discrete case. Thus, the techniques employed could be useful for lifting other discrete coding theorems to the case of arbitrary alphabets.

II. MAIN RESULT

Let Y and X be random variables with arbitrary alphabets \mathcal{S}_Y and \mathcal{S}_X , respectively. The bold-faced random vectors \mathbf{Y} and \mathbf{X} are n i.i.d. copies of Y and X , respectively. We then have the following definitions.

Definition 1. A pair $(S, R) \in \mathbb{R}^2$ is achievable if for some $n \in \mathbb{N}$ there exists a measurable function $f: \mathcal{S}_X^n \rightarrow \mathcal{M}$ for some finite set \mathcal{M} with bounded cardinality $\frac{1}{n} \log|\mathcal{M}| \leq R$ and

$$\frac{1}{n} I(\mathbf{Y}; f(\mathbf{X})) \geq S. \quad (2)$$

The set of all achievable pairs is denoted $\mathcal{R} \subseteq \mathbb{R}^2$.

Definition 2. A pair $(S, R) \in \mathbb{R}^2$ is IB-achievable if there exists an additional random variable U with arbitrary alphabet \mathcal{S}_U , satisfying $Y \dashv\vdash X \dashv\vdash U$ and

$$R \geq I(X; U), \quad (3)$$

$$S \leq I(Y; U). \quad (4)$$

The set of all IB-achievable pairs is denoted $\mathcal{R}_{\text{IB}} \subseteq \mathbb{R}^2$.

In what follows, we will prove the following theorem.

Theorem 3. The equality $\overline{\mathcal{R}_{\text{IB}}} = \overline{\mathcal{R}}$ holds.

III. PRELIMINARIES

When introducing a function, we implicitly assume it to be measurable w.r.t. the appropriate σ -algebras. The σ -algebra associated with a finite set is its power set and the σ -algebra associated with \mathbb{R} is the Borel σ -algebra. The symbol \emptyset is used for the empty set and for a constant random variable. When there is no possibility for confusion, we will not distinguish between a single-element set and its element, e. g., we write x instead of $\{x\}$ and $\mathbb{1}_x$ for the indicator function of $\{x\}$. We use $A \Delta B := (A \setminus B) \cup (B \setminus A)$ to denote the symmetric set difference.

Let (Ω, Σ, μ) be a probability space. A random variable $X: \Omega \rightarrow \mathcal{S}_X$ takes values in the measurable space $(\mathcal{S}_X, \mathcal{A}_X)$. The push-forward probability measure $\mu_X: \mathcal{A}_X \rightarrow [0, 1]$ is defined by $\mu_X(A) = \mu(X^{-1}(A))$ for all $A \in \mathcal{A}_X$. We will state most results in terms of push-forward measures and usually ignore the background probability space. When multiple random variables are defined, we implicitly assume the push-forward measures to be consistent in the sense that, e. g., $\mu_X(A) = \mu_{XY}(A \times \mathcal{S}_Y)$ for all $A \in \mathcal{A}_X$.

For $n \in \mathbb{N}$ let Ω^n denote the n -fold Cartesian product of (Ω, Σ, μ) . A bold-faced random vector, e. g., \mathbf{X} , defined on Ω^n , is an n -fold copy of X , i. e., $\mathbf{X} = X^n$. Accordingly, the corresponding push-forward measure, e. g., $\mu_{\mathbf{X}}$ is the n -fold product measure.

For a random variable X let a_X , b_X , and c_X denote arbitrary functions on \mathcal{S}_X , each with finite range. We will use the symbol \mathcal{M}_X to denote the range of a_X , i. e., $a_X: \mathcal{S}_X \rightarrow \mathcal{M}_X$.

Definition 4 ([18, Def. 8.11]). *The conditional expectation of a random variable X with $\mathcal{S}_X = \mathbb{R}$, given a random variable Y , is a random variable $\mathbb{E}[X|Y]$ such that*

- 1) $\mathbb{E}[X|Y]$ is $\sigma(Y)$ -measurable, and
- 2) for all $A \in \sigma(Y)$, we have $\mathbb{E}[\mathbb{1}_A \mathbb{E}[X|Y]] = \mathbb{E}[\mathbb{1}_A X]$.

The conditional probability of an event $B \in \Sigma$ given Y is defined as $P\{B|Y\} := \mathbb{E}[\mathbb{1}_B|Y]$.

The conditional expectation and therefore also the conditional probability exists and is unique up to equality almost surely by [18, Thm. 8.12]. Furthermore, if $(\mathcal{S}_X, \mathcal{A}_X)$ is a standard space [17, Sec. 1.5], there even exists a *regular conditional distribution* of X given Y [18, Thm. 8.37].

Definition 5. *For two random variables X and Y a regular conditional distribution of X given Y is a function $\kappa_{X|Y}: \Omega \times \mathcal{A}_X \rightarrow [0, 1]$ such that*

- 1) for every $\omega \in \Omega$, the set function $\kappa_{X|Y}(\omega) := \kappa_{X|Y}(\omega; \cdot)$ is a probability measure on $(\mathcal{S}_X, \mathcal{A}_X)$.
- 2) for every set $A \in \mathcal{A}_X$, the function $\kappa_{X|Y}(\cdot; A)$ is $\sigma(Y)$ -measurable.
- 3) for μ -a. e. $\omega \in \Omega$ and all $A \in \mathcal{A}_X$, we have $\kappa_{X|Y}(\omega; A) = P\{X^{-1}(A)|Y\}(\omega)$ (cf. Def. 4).

Note, in particular, that finite spaces are standard spaces.

Remark 1. If the random variable Y is discrete, then $\kappa_{X|Y}$ reduces to conditioning given events $Y = y$ for $y \in \mathcal{S}_Y$, i. e., $\kappa_{X|Y}(\omega; A) = \frac{\mu_{XY}(A \times Y(\omega))}{\mu_Y(Y(\omega))}$ (cf. [18, Lem. 8.10]).

We use the following definitions and results from [17], [18].

Definition 6. *For random variables X and Y with $|\mathcal{S}_X| < \infty$ the conditional entropy is defined as [17, Sec. 5.5]*

$$H(X|Y) := \int H(\kappa_{X|Y}) d\mu, \quad (5)$$

where $H(\cdot)$ denotes discrete entropy on \mathcal{S}_X . For arbitrary random variables X , Y , and Z the conditional mutual information is defined as [17, Lem. 5.5.7]

$$\begin{aligned} I(X; Y|Z) &:= \sup_{\alpha_X, \alpha_Y} \int D(\kappa_{\alpha_X(X)\alpha_Y(Y)|Z} \| \kappa_{\alpha_X(X)|Z} \times \kappa_{\alpha_Y(Y)|Z}) d\mu \quad (6) \\ &= \sup_{\alpha_X, \alpha_Y} [H(\alpha_X(X)|Z) + H(\alpha_Y(Y)|Z) - H(\alpha_X(X)\alpha_Y(Y)|Z)], \quad (7) \end{aligned}$$

where $D(\cdot \| \cdot)$ denotes Kullback-Leibler divergence [17, Sec. 2.3] and the supremum is taken over all α_X and α_Y with finite range. The mutual information is given by [17, Lem. 5.5.1] $I(X; Y) := I(X; Y|\emptyset)$.

Definition 7 ([18, Def. 12.20]). *For arbitrary random variables X , Y , and Z , the Markov chain $X \dashv\vdash Y \dashv\vdash Z$ holds if, for any $A \in \mathcal{A}_X$, $B \in \mathcal{A}_Z$, the following holds μ -a. e.:*

$$P\{X^{-1}(A) \cap Z^{-1}(B)|Y\} = P\{X^{-1}(A)|Y\}P\{Z^{-1}(B)|Y\}. \quad (8)$$

In the following, we collect some properties of these definitions.

Lemma 8. *For random variables X , Y , and Z the following properties hold:*

- (i) $I(X; Y|Z) \geq 0$ with equality if and only if $X \dashv\vdash Z \dashv\vdash Y$.
- (ii) For discrete X , i. e., $|\mathcal{S}_X| < \infty$, we have $I(X; Y) = H(X) - H(X|Y)$.
- (iii) $I(X; YZ) = I(X; Z) + I(X; Y|Z)$.
- (iv) If $X \dashv\vdash Y \dashv\vdash Z$, then $I(X; Y) \geq I(X; Z)$.

Proof. (i): The claim $I(X; Y|Z) \geq 0$ follows directly from (6) and the non-negativity of divergence.

Assume that $X \dashv\vdash Z \dashv\vdash Y$, i. e., $P\{X^{-1}(A) \cap Y^{-1}(B)|Z\} = P\{X^{-1}(A)|Z\}P\{Y^{-1}(B)|Z\}$ almost everywhere. Let $\alpha_X: \mathcal{S}_X \rightarrow \mathcal{M}_X$ and $\alpha_Y: \mathcal{S}_Y \rightarrow \mathcal{M}_Y$ be functions with finite range. Pick two arbitrary sets $A \subseteq \mathcal{M}_X$, $B \subseteq \mathcal{M}_Y$ and we obtain μ -a. e.

$$\begin{aligned} \kappa_{\alpha_X(X)\alpha_Y(Y)|Z}(\cdot; A \times B) &= P\{X^{-1}(\alpha_X^{-1}(A)) \cap Y^{-1}(\alpha_Y^{-1}(B))|Z\} \quad (9) \\ &= P\{X^{-1}(\alpha_X^{-1}(A))|Z\}P\{Y^{-1}(\alpha_Y^{-1}(B))|Z\} \quad (10) \\ &= \kappa_{\alpha_X(X)|Z}(\cdot; A)\kappa_{\alpha_Y(Y)|Z}(\cdot; B), \quad (11) \end{aligned}$$

where (9) and (11) follow from part 3 of Def. 5. This proves that μ -a. e. the equality of measures $\kappa_{\alpha_X(X)\alpha_Y(Y)|Z} = \kappa_{\alpha_X(X)|Z} \times \kappa_{\alpha_Y(Y)|Z}$ holds. By the properties of Kullback-Leibler divergence [17, Thm. 2.3.1] we have $I(X; Y|Z) = 0$ due to (6).

On the other hand, assume $I(X; Y|Z) = 0$ and choose arbitrary sets $A \in \mathcal{A}_X$ and $B \in \mathcal{A}_Y$. We define $\alpha_X := \mathbb{1}_A$, $\alpha_Y := \mathbb{1}_B$, $\hat{X} := \alpha_X(X)$, and $\hat{Y} := \alpha_Y(Y)$. By (6) we have $D(\kappa_{\hat{X}\hat{Y}|Z}(\omega) \| \kappa_{\hat{X}|Z}(\omega) \times \kappa_{\hat{Y}|Z}(\omega)) = 0$ for μ -a. e. $\omega \in \Omega$, which is equivalent to the equality μ -a. e. of the measures $\kappa_{\hat{X}\hat{Y}|Z} = \kappa_{\hat{X}|Z} \times \kappa_{\hat{Y}|Z}$. We obtain μ -a. e.,

$$P\{X^{-1}(A) \cap Y^{-1}(B)|Z\} = \kappa_{\hat{X}\hat{Y}|Z}(\cdot; 1 \times 1) \quad (12)$$

$$= \kappa_{\hat{X}|Z}(\cdot; 1)\kappa_{\hat{Y}|Z}(\cdot; 1) \quad (13)$$

$$= P\{X^{-1}(A)|Z\}P\{Y^{-1}(B)|Z\}. \quad (14)$$

(ii): See [17, Lem. 5.5.6].

(iii): See [17, Lem. 5.5.7].

(iv): Using Prop. (i) we have $I(X; Z|Y) = 0$ and by Prop. (iii) it follows that

$$\begin{aligned} I(X; Z) &\leq I(X; YZ) \quad (15) \\ &= I(X; Y) + I(X; Z|Y) = I(X; Y). \quad \blacksquare \end{aligned}$$

Occasionally we will interpret a probability measure on a finite space \mathcal{M} as a vector in $[0, 1]^{\mathcal{M}}$, equipped with the Borel σ -algebra. We will use the L_∞ -distance on this space.

Definition 9. *For two probability measures μ and ν on a finite space \mathcal{M} , their distance is defined as the L_∞ -distance $d(\mu, \nu) := \max_{m \in \mathcal{M}} |\mu(m) - \nu(m)|$. The diameter of $A \subseteq [0, 1]^{\mathcal{M}}$ is defined as $\text{diam}(A) = \sup_{\mu, \nu \in A} d(\mu, \nu)$.*

Lemma 10 ([19, Lem. 2.7]). *For two probability measures μ and ν on a finite space \mathcal{M} with $d(\mu, \nu) \leq \varepsilon \leq \frac{1}{2}$ the inequality $|H(\mu) - H(\nu)| \leq -\varepsilon|\mathcal{M}|\log \varepsilon$ holds.*

IV. PROOF OF $\mathcal{R}_{\text{IB}} \subseteq \overline{\mathcal{R}}$

For finite spaces \mathcal{S}_Y , \mathcal{S}_X , and \mathcal{S}_U , the statement $\mathcal{R}_{\text{IB}} \subseteq \overline{\mathcal{R}}$ is well known, cf., [9, Sec. IV], [10, Sec. III.F]. We restate it in the form of the following lemma.

Lemma 11. *For random variables Y , X , and U with finite \mathcal{S}_Y , \mathcal{S}_X , and \mathcal{S}_U , assume that $Y \dashv\vdash X \dashv\vdash U$ holds. Then, for any $\varepsilon > 0$, there exists $n \in \mathbb{N}$ and a function $f: \mathcal{S}_X^n \rightarrow \mathcal{M}$ with $\frac{1}{n} \log |\mathcal{M}| \leq I(X; U) + \varepsilon$ such that $\frac{1}{n} I(\mathbf{Y}; f(\mathbf{X})) \geq I(Y; U) - \varepsilon$.*

In a first step, we will utilize Lem. 11 to show $\mathcal{R}_{\text{IB}} \subseteq \overline{\mathcal{R}}$ for an arbitrary alphabet \mathcal{S}_X , i. e., we wish to prove the following Proposition 12, lifting the restriction $|\mathcal{S}_X| < \infty$.

Proposition 12. *For random variables Y , X , and U with finite \mathcal{S}_Y and \mathcal{S}_U , assume that $Y \dashv\vdash X \dashv\vdash U$ holds. Then, for any*

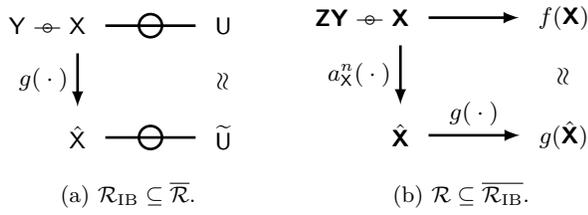


Fig. 1: Illustrations.

$\varepsilon > 0$, there exists $n \in \mathbb{N}$ and a function $f: \mathcal{S}_X^n \rightarrow \mathcal{M}$ with $\frac{1}{n} \log |\mathcal{M}| \leq I(\mathbf{X}; \mathbf{U}) + \varepsilon$ such that

$$\frac{1}{n} I(\mathbf{Y}; f(\mathbf{X})) \geq I(\mathbf{Y}; \mathbf{U}) - \varepsilon. \quad (16)$$

Remark 2. Considering that both definitions of achievability (Defs. 1 and 2) only rely on the notion of mutual information, one may assume that Def. 6 can be used to directly infer Proposition 12 from Lem. 11. However, this is not the case. For an arbitrary discretization $a_X(\mathbf{X})$ of \mathbf{X} , we do have $I(a_X(\mathbf{X}); \mathbf{U}) \leq I(\mathbf{X}; \mathbf{U})$. However, the Markov chain $\mathbf{Y} \rightarrow a_X(\mathbf{X}) \rightarrow \mathbf{U}$ does not hold in general. To circumvent this problem, we will use a discrete random variable $\hat{\mathbf{X}} = g(\mathbf{X})$ with an appropriate quantizer g and construct a new random variable $\tilde{\mathbf{U}}$, satisfying the Markov chain $\mathbf{Y} \rightarrow \hat{\mathbf{X}} \rightarrow \tilde{\mathbf{U}}$ such that $I(\mathbf{Y}; \tilde{\mathbf{U}})$ is close to $I(\mathbf{Y}; \mathbf{U})$. Fig. 1a illustrates this strategy. We choose the quantizer g based on the conditional probability distribution of \mathbf{U} given \mathbf{X} , i. e., quantization based on $\kappa_{\mathbf{U}|\mathbf{X}}$ using L_∞ -distance (cf. Def. 9). Subsequently, we will use that, by Lem. 10, a small L_∞ -distance guarantees a small gap in terms of information measures.

Proof of Proposition 12. Let $\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}}$ be a probability measure on $\Omega := \mathcal{S}_Y \times \mathcal{S}_X \times \mathcal{S}_U$, such that $\mathbf{Y} \rightarrow \mathbf{X} \rightarrow \mathbf{U}$ holds. Fix $0 < \delta \leq \frac{1}{2}$ and find a finite, measurable partition $(P_i)_{i \in \mathcal{I}}$ of the space of probability measures on \mathcal{S}_U such that for every $i \in \mathcal{I}$ we have $\text{diam}(P_i) \leq \delta$ and fix some $\nu_i \in P_i$ for every $i \in \mathcal{I}$. Define the random variable $\hat{\mathbf{X}}: \Omega \rightarrow \mathcal{I}$ as $\hat{\mathbf{X}} = i$ if $\kappa_{\mathbf{U}|\mathbf{X}} \in P_i$. The random variable $\hat{\mathbf{X}}$ is $\sigma(\mathbf{X})$ -measurable (see Appendix A). We can therefore find a measurable function g such that $\hat{\mathbf{X}} = g(\mathbf{X})$ by the factorization lemma [18, Corollary 1.97]. Define the new probability space $\Omega \times \prod_{i \in \mathcal{I}} \mathcal{S}_U$, equipped with the probability measure $\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}\tilde{\mathbf{U}}_Z} := \mu_{\mathbf{Y}\mathbf{X}\mathbf{U}} \times \prod_{i \in \mathcal{I}} \nu_i$. Slightly abusing notation, we define the random variables \mathbf{Y} , \mathbf{X} , \mathbf{U} , and $\tilde{\mathbf{U}}_i$ (for every $i \in \mathcal{I}$) as the according projections. We also use $\hat{\mathbf{X}} = g(\mathbf{X})$ and define the random variable $\tilde{\mathbf{U}} = \tilde{\mathbf{U}}_{\hat{\mathbf{X}}}$. From this construction we have $\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}\tilde{\mathbf{U}}_Z}$ -a. e. the equality of measures $\kappa_{\tilde{\mathbf{U}}|\hat{\mathbf{X}}} = \kappa_{\mathbf{U}|\mathbf{X}} = \nu_{\hat{\mathbf{X}}}$, as well as $\mathbf{Y} \rightarrow \hat{\mathbf{X}} \rightarrow \tilde{\mathbf{U}}$ and $\mathbf{Y} \rightarrow \mathbf{X} \rightarrow \tilde{\mathbf{U}}$. This is proven in the extended version [20]. Therefore, we have $\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}\tilde{\mathbf{U}}_Z}$ -a. e.

$$d(\kappa_{\tilde{\mathbf{U}}|\hat{\mathbf{X}}}, \kappa_{\mathbf{U}|\mathbf{X}}) \leq \delta, \quad \text{and} \quad d(\kappa_{\tilde{\mathbf{U}}|\hat{\mathbf{X}}}, \kappa_{\mathbf{U}|\mathbf{X}}) \leq \delta, \quad (17)$$

by $\kappa_{\tilde{\mathbf{U}}|\hat{\mathbf{X}}} = \kappa_{\mathbf{U}|\mathbf{X}} = \nu_{\hat{\mathbf{X}}}$ and $\kappa_{\mathbf{U}|\mathbf{X}}, \nu_{\hat{\mathbf{X}}} \in P_{\hat{\mathbf{X}}}$. Thus, for any $u \in \mathcal{S}_U$,

$$\mu_{\mathbf{U}}(u) = \int \kappa_{\mathbf{U}|\mathbf{X}}(\cdot; u) d\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}} \quad (18)$$

$$\leq \int (\kappa_{\tilde{\mathbf{U}}|\hat{\mathbf{X}}}(\cdot; u) + \delta) d\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}\tilde{\mathbf{U}}_Z} = \mu_{\tilde{\mathbf{U}}}^{\sim}(u) + \delta \quad (19)$$

and, by the same argument, $\mu_{\mathbf{U}}(u) \geq \mu_{\tilde{\mathbf{U}}}^{\sim}(u) - \delta$, i. e., in total,

$$d(\mu_{\mathbf{U}}, \mu_{\tilde{\mathbf{U}}}^{\sim}) \leq \delta. \quad (20)$$

Thus, we obtain

$$I(\mathbf{X}; \mathbf{U}) = H(\mu_{\mathbf{U}}) - H(\mathbf{U}|\mathbf{X}) \quad (21)$$

$$\stackrel{(20)}{\geq} H(\mu_{\tilde{\mathbf{U}}}^{\sim}) + \delta |\mathcal{S}_U| \log \delta - \int H(\kappa_{\mathbf{U}|\mathbf{X}}) d\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}} \quad (22)$$

$$\stackrel{(17)}{\geq} H(\mu_{\tilde{\mathbf{U}}}^{\sim}) + 2\delta |\mathcal{S}_U| \log \delta - \int H(\kappa_{\tilde{\mathbf{U}}|\hat{\mathbf{X}}}) d\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}\tilde{\mathbf{U}}_Z} \quad (23)$$

$$= I(\hat{\mathbf{X}}; \tilde{\mathbf{U}}) + 2\delta |\mathcal{S}_U| \log \delta, \quad (24)$$

where (21) and (24) follow from Prop. (ii) of Lem. 8, and in both (22) and (23) we used Lem. 10. From $\mathbf{Y} \rightarrow \mathbf{X} \rightarrow \mathbf{U}$ and Prop. (i) of Lem. 8, we know that $\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}}$ -a. e., we have the equality of measures $\kappa_{\mathbf{Y}\mathbf{U}|\mathbf{X}} = \kappa_{\mathbf{Y}|\mathbf{X}} \times \kappa_{\mathbf{U}|\mathbf{X}}$. Using this equality in (26) we obtain

$$\mu_{\mathbf{Y}\mathbf{U}}(y \times u) = \int \kappa_{\mathbf{Y}\mathbf{U}|\mathbf{X}}(\cdot; y \times u) d\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}} \quad (25)$$

$$= \int \kappa_{\mathbf{Y}|\mathbf{X}}(\cdot; y) \kappa_{\mathbf{U}|\mathbf{X}}(\cdot; u) d\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}} \quad (26)$$

$$\stackrel{(17)}{\leq} \int \kappa_{\mathbf{Y}|\mathbf{X}}(\cdot; y) (\kappa_{\tilde{\mathbf{U}}|\hat{\mathbf{X}}}(\cdot; u) + \delta) d\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}\tilde{\mathbf{U}}_Z} \quad (27)$$

$$\leq \int \kappa_{\tilde{\mathbf{Y}}|\hat{\mathbf{X}}}(\cdot; y \times u) d\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}\tilde{\mathbf{U}}_Z} + \delta \quad (28)$$

$$= \mu_{\tilde{\mathbf{Y}}\tilde{\mathbf{U}}}^{\sim}(y \times u) + \delta, \quad (29)$$

where (25) and (29) follow from the defining property of conditional probability, part 2 of Def. 4, and (28) follows from $\mathbf{Y} \rightarrow \mathbf{X} \rightarrow \tilde{\mathbf{U}}$ and Prop. (i) of Lem. 8. By the same argument, one can show that $\mu_{\mathbf{Y}\mathbf{U}}(y \times u) \geq \mu_{\tilde{\mathbf{Y}}\tilde{\mathbf{U}}}^{\sim}(y \times u) - \delta$. Therefore, in total, $d(\mu_{\mathbf{Y}\mathbf{U}}, \mu_{\tilde{\mathbf{Y}}\tilde{\mathbf{U}}}^{\sim}) \leq \delta$ and, by Lem. 10,

$$|H(\mathbf{Y}\mathbf{U}) - H(\mathbf{Y}\tilde{\mathbf{U}})| \leq -\delta |\mathcal{S}_Y| |\mathcal{S}_U| \log \delta. \quad (30)$$

Thus, the mutual information can be bounded by

$$I(\mathbf{Y}; \mathbf{U}) = H(\mathbf{Y}) + H(\mathbf{U}) - H(\mathbf{Y}\mathbf{U}) \quad (31)$$

$$\stackrel{(20)}{\leq} H(\mathbf{Y}) + H(\tilde{\mathbf{U}}) - \delta |\mathcal{S}_U| \log \delta - H(\mathbf{Y}\mathbf{U}) \quad (32)$$

$$\stackrel{(30)}{\leq} I(\mathbf{Y}; \tilde{\mathbf{U}}) - \delta (|\mathcal{S}_Y| + 1) |\mathcal{S}_U| \log \delta \quad (33)$$

$$\leq I(\mathbf{Y}; \tilde{\mathbf{U}}) - 2\delta |\mathcal{S}_Y| |\mathcal{S}_U| \log \delta, \quad (34)$$

where we applied Lem. 10 in (32) and (33). We apply Lem. 11 to the three random variables \mathbf{Y} , $\hat{\mathbf{X}}$, and $\tilde{\mathbf{U}}$ and obtain a function $\hat{f}: \mathcal{I}^n \rightarrow \mathcal{M}$ with $\frac{1}{n} I(\mathbf{Y}; \hat{f}(\hat{\mathbf{X}})) \geq I(\mathbf{Y}; \tilde{\mathbf{U}}) - \delta$ and

$$\frac{1}{n} \log |\mathcal{M}| \leq I(\hat{\mathbf{X}}; \tilde{\mathbf{U}}) + \delta \stackrel{(24)}{\leq} I(\mathbf{X}; \mathbf{U}) + \delta - 2\delta |\mathcal{S}_U| \log \delta. \quad (35)$$

We have $\hat{\mathbf{X}} = g^n \circ \mathbf{X}$ and defining $f := \hat{f} \circ g^n$, we obtain

$$\frac{1}{n} I(\mathbf{Y}; f(\mathbf{X})) = \frac{1}{n} I(\mathbf{Y}; \hat{f}(\hat{\mathbf{X}})) \geq I(\mathbf{Y}; \tilde{\mathbf{U}}) - \delta \quad (36)$$

$$\stackrel{(34)}{\geq} I(\mathbf{Y}; \mathbf{U}) + 2\delta |\mathcal{S}_Y| |\mathcal{S}_U| \log \delta - \delta. \quad (37)$$

Choosing δ such that $\varepsilon \geq -2\delta |\mathcal{S}_Y| |\mathcal{S}_U| \log \delta + \delta$ completes the proof. \blacksquare

We can now complete the proof by showing the following lemma.

Lemma 13. $\mathcal{R}_{\text{IB}} \subseteq \bar{\mathcal{R}}$.

Proof. Assuming $(S, R) \in \mathcal{R}_{\text{IB}}$, choose $\mu_{\mathbf{Y}\mathbf{X}\mathbf{U}}$ according to Def. 2. Clearly $I(\mathbf{X}; \mathbf{U}) < \infty$ to satisfy (3) and thus also $I(\mathbf{Y}; \mathbf{U}) < \infty$ by Prop. (iv) of Lem. 8 as $\mathbf{Y} \rightarrow \mathbf{X} \rightarrow \mathbf{U}$ holds. Pick $\varepsilon > 0$, select

functions a_X, a_U such that $I(a_X(X); a_U(U)) \geq I(X; U) - \varepsilon$, and select functions b_Y, b_U such that $I(b_Y(Y); b_U(U)) \geq I(Y; U) - \varepsilon$ (cf. (7)). Using $\hat{U} := (a_U(U), b_U(U))$ and $\hat{Y} := b_Y(Y)$, we have

$$0 = I(Y; U|X) = \sup_{c_Y, c_U} I(c_Y(Y); c_U(U)|X) \geq I(\hat{Y}; \hat{U}|X) \geq 0 \quad (38)$$

as well as

$$I(X; U) = \sup_{c_X, c_U} I(c_X(X); c_U(U)) \quad (39)$$

$$\geq \sup_{c_X} I(c_X(X); \hat{U}) = I(X; \hat{U}), \quad \text{and} \quad (40)$$

$$I(Y; U) - \varepsilon \leq I(b_Y(Y); b_U(U)) \leq I(\hat{Y}; \hat{U}). \quad (41)$$

We apply Proposition 12, substituting $\hat{U} \rightarrow U$ and $\hat{Y} \rightarrow Y$. Proposition 12 guarantees the existence of a function $f: S_X^n \rightarrow \mathcal{M}$ with $\frac{1}{n} \log |\mathcal{M}| \leq I(X; \hat{U}) + \varepsilon \leq I(X; U) + \varepsilon \leq R + \varepsilon$ and

$$\frac{1}{n} I(\mathbf{Y}; f(\mathbf{X})) = \frac{1}{n} \sup_{c_Y} I(c_Y \circ \mathbf{Y}; f(\mathbf{X})) \quad (42)$$

$$\geq \frac{1}{n} I(b_Y^n \circ \mathbf{Y}; f(\mathbf{X})) = \frac{1}{n} I(\hat{\mathbf{Y}}; f(\mathbf{X})) \quad (43)$$

$$\stackrel{(16)}{\geq} I(\hat{\mathbf{Y}}; \hat{U}) - \varepsilon \stackrel{(41)}{\geq} I(Y; U) - 2\varepsilon \stackrel{(4)}{\geq} S - 2\varepsilon. \quad (44)$$

Thus, $(S - 2\varepsilon, R - \varepsilon) \in \mathcal{R}$ and therefore $(S, R) \in \overline{\mathcal{R}}$. ■

V. PROOF OF $\mathcal{R} \subseteq \overline{\mathcal{R}}_{\text{IB}}$

We start with the well-known result $\mathcal{R}_{\text{IB}} \subseteq \overline{\mathcal{R}}$ for finite spaces S_Y, S_X , and S_U , cf., [9, Sec. IV], [10, Sec. III.F]. The statement is rephrased in the following lemma.

Lemma 14. *Assume that the spaces S_Y and S_X are both finite and μ_{YX} is fixed. For some $n \in \mathbb{N}$, let $f: S_X^n \rightarrow \mathcal{M}$ be a function with $|\mathcal{M}| < \infty$. Then there exists a probability measure μ_{YXU} , extending μ_{YX} , such that S_U is finite, $Y \circlearrowleft X \circlearrowleft U$, and*

$$I(X; U) \leq \frac{1}{n} \log |\mathcal{M}|, \quad (45)$$

$$I(Y; U) \geq \frac{1}{n} I(\mathbf{Y}; f(\mathbf{X})). \quad (46)$$

We can slightly strengthen Lem. 14.

Corollary 15. *Assume that, in the setting of Lem. 14, we are given μ_{ZYX} on $S_Z \times S_Y \times S_X$, extending μ_{YX} , where S_Z is arbitrary, not necessarily finite. Then there exists a probability measure μ_{ZYXU} , extending μ_{ZYX} , such that S_U is finite and $ZY \circlearrowleft X \circlearrowleft U$, (45), and (46) hold.*

Proof. Apply Lem. 14 to obtain μ_{YXU} on $S_Y \times S_X \times S_U$ satisfying (45), (46), and $Y \circlearrowleft X \circlearrowleft U$. We define μ_{ZYXU} by

$$\mu_{ZYXU}(A \times y \times x \times u) = \frac{\mu_{ZYX}(A \times y \times x)}{\mu_{YX}(y \times x)} \mu_{YXU}(y \times x \times u) \quad (47)$$

for any $(y, x, u) \in S_Y \times S_X \times S_U$ and $A \in \mathcal{A}_Z$. Pick arbitrary $A \in \mathcal{A}_Z$, $y \in S_Y$, and $u \in S_U$. The Markov chain $ZY \circlearrowleft X \circlearrowleft U$ now follows as the events $Z^{-1}(A) \cap Y^{-1}(y)$ and $U^{-1}(u)$ are independent given $X^{-1}(x)$ for any $x \in S_X$ (cf. Rmk. 1). ■

Again, we proceed by extending Cor. 15, lifting the restriction that S_X is finite and obtain the following proposition.

Proposition 16. *Given a probability measure μ_{ZYX} as in Cor. 15, assume that $|\mathcal{S}_Y| < \infty$. For some $n \in \mathbb{N}$, let $f: S_X^n \rightarrow \mathcal{M}$ be a function with $|\mathcal{M}| < \infty$. Then, for any $\varepsilon > 0$, there exists a*

probability measure μ_{ZYXU} , extending μ_{ZYX} with $ZY \circlearrowleft X \circlearrowleft U$ and

$$I(X; U) \leq \frac{1}{n} \log |\mathcal{M}| \quad (48)$$

$$I(Y; U) \geq \frac{1}{n} I(\mathbf{Y}; f(\mathbf{X})) - \varepsilon. \quad (49)$$

Remark 3. In contrast to Proposition 12, Proposition 16 could be proved by the usual single-letterization + time-sharing strategy, by showing that the necessary Markov chains hold. However, we will rely on the discrete case (Lem. 14) and showcase a technique to lift it to general alphabets.

Remark 4. In the proof of Proposition 16, we face a similar problem as outlined in Rmk. 2. We need to construct a function $g(\hat{\mathbf{X}})$ of a “per-letter” quantization $\hat{\mathbf{X}} := a_X^n(\mathbf{X})$, that is close to $f(\mathbf{X})$ in distribution. Fig. 1b provides a sketch.

Proof of Proposition 16. We can partition $S_X^n = \bigcup_{m \in \mathcal{M}} \mathcal{Q}_m$ into finitely many measurable, mutually disjoint sets $\mathcal{Q}_m := f^{-1}(m)$, $m \in \mathcal{M}$. We want to approximate the sets \mathcal{Q}_m by a finite union of rectangles in the semiring [18, Def. 1.9] $\Xi := \{\mathcal{B} : \mathcal{B} = \bigtimes_{i=1}^n B_i \text{ with } B_i \in \mathcal{A}_X\}$. We choose $\delta > 0$, which will be specified later. According to [18, Thm. 1.65(ii)], we obtain $\mathcal{B}^{(m)} := \bigcup_{k=1}^K \mathcal{B}_k^{(m)}$ for each $m \in \mathcal{M}$, where $\mathcal{B}_k^{(m)} \in \Xi$ are mutually disjoint sets, satisfying $\mu_X(\mathcal{B}^{(m)} \Delta \mathcal{Q}_m) \leq \delta$. Since $\mathcal{B}_k^{(m)} \in \Xi$, we have $\mathcal{B}_k^{(m)} = \bigtimes_{i=1}^n B_{k,i}^{(m)}$ for some $B_{k,i}^{(m)} \in \mathcal{A}_X$. We can construct functions a_X and g such that $g \circ a_X^n(\mathbf{x}) = m$ whenever $\mathbf{x} \in \mathcal{B}^{(m)}$ and $\mathbf{x} \notin \mathcal{B}^{(m')}$ with $\mathcal{B}^{(m')} := \bigcup_{m' \neq m} \mathcal{B}^{(m')}$. Indeed, we obtain a_X by finding a measurable partition of S_X that is finer than $(B_{k,i}^{(m)}, (B_{k,i}^{(m)})^c)$ for all i, k, m . For fixed $m \in \mathcal{M}$,

$$\mathcal{Q}_m \subseteq \mathcal{Q}_m \cup (\mathcal{B}^{(m)} \setminus \mathcal{B}^{(m)}) \quad (50)$$

$$\subseteq (\mathcal{B}^{(m)} \setminus \mathcal{B}^{(m)}) \cup (\mathcal{Q}_m \setminus \mathcal{B}^{(m)}) \cup \bigcup_{m' \neq m} \mathcal{Q}_m \cap \mathcal{B}^{(m')} \quad (51)$$

$$\subseteq (\mathcal{B}^{(m)} \setminus \mathcal{B}^{(m)}) \cup (\mathcal{Q}_m \Delta \mathcal{B}^{(m)}) \cup \bigcup_{m' \neq m} \mathcal{B}^{(m')} \setminus \mathcal{Q}_{m'} \quad (52)$$

$$\subseteq (\mathcal{B}^{(m)} \setminus \mathcal{B}^{(m)}) \cup \bigcup_{m'} \mathcal{B}^{(m')} \Delta \mathcal{Q}_{m'}, \quad (53)$$

where we used the fact that $\mathcal{Q}_m \cap \mathcal{Q}_{m'} = \emptyset$ for $m \neq m'$ in (52). Using $\hat{X} := a_X(X)$, we obtain for any $\mathbf{y} \in S_Y^n$

$$\mu_{Yf(\mathbf{X})}(\mathbf{y} \times m) = \mu_{YX}(\mathbf{y} \times \mathcal{Q}_m) \quad (54)$$

$$\stackrel{(53)}{\leq} \mu_{YX}(\mathbf{y} \times (\mathcal{B}^{(m)} \setminus \mathcal{B}^{(m)})) + \sum_{m'} \mu_X(\mathcal{B}^{(m')} \Delta \mathcal{Q}_{m'}) \quad (55)$$

$$\leq \mu_{Yg(\hat{X})}(\mathbf{y} \times m) + |\mathcal{M}| \delta. \quad (56)$$

On the other hand, we have

$$\mu_{Yf(\mathbf{X})}(\mathbf{y} \times m) = \mu_Y(\mathbf{y}) - \sum_{m' \neq m} \mu_{Yf(\mathbf{X})}(\mathbf{y} \times m') \quad (57)$$

$$\stackrel{(56)}{\geq} \mu_Y(\mathbf{y}) - \sum_{m' \neq m} (\mu_{Yg(\hat{X})}(\mathbf{y} \times m') + |\mathcal{M}| \delta) \quad (58)$$

$$\geq \mu_{Yg(\hat{X})}(\mathbf{y} \times m) - |\mathcal{M}|^2 \delta. \quad (59)$$

We thus obtain $d(\mu_{Yf(\mathbf{X})}, \mu_{Yg(\hat{X})}) \leq |\mathcal{M}|^2 \delta$. This also implies $d(\mu_{f(\mathbf{X})}, \mu_{g(\hat{X})}) \leq |\mathcal{S}_Y|^n |\mathcal{M}|^2 \delta$. Assume $|\mathcal{S}_Y|^n |\mathcal{M}|^2 \delta \leq \frac{1}{2}$ and apply Cor. 15 substituting $\hat{X} \rightarrow X, XZ \rightarrow Z$, and the function $g \rightarrow f$. This yields a random variable U with $XZY \circlearrowleft \hat{X} \circlearrowleft U$,

$$I(\hat{X}; U) \leq \frac{1}{n} \log |\mathcal{M}|, \quad \text{and} \quad I(Y; U) \geq \frac{1}{n} I(\mathbf{Y}; g(\hat{\mathbf{X}})). \quad (60)$$

We also obtain $\mathbf{ZY} \dashv\!\!\!\dashv \mathbf{X} \dashv\!\!\!\dashv \mathbf{U}$ due to

$$0 = I(\mathbf{XZY}; \mathbf{U}|\hat{\mathbf{X}}) \quad (61)$$

$$= I(\mathbf{XZY}; \mathbf{U}) - I(\mathbf{U}; \hat{\mathbf{X}}) \quad (62)$$

$$\geq I(\mathbf{XZY}; \mathbf{U}) - I(\mathbf{U}; \mathbf{X}) \quad (63)$$

$$= I(\mathbf{ZY}; \mathbf{U}|\mathbf{X}) \quad (64)$$

$$\geq 0, \quad (65)$$

where (61) follows from $\mathbf{XZY} \dashv\!\!\!\dashv \hat{\mathbf{X}} \dashv\!\!\!\dashv \mathbf{U}$ using Prop. (i) of Lem. 8, (62) and (64) follow from Prop. (iii) of Lem. 8, (63) is a consequence of Def. 6, and we used Prop. (i) of Lem. 8 in (65). This also immediately implies $0 = I(\mathbf{X}; \mathbf{U}|\hat{\mathbf{X}})$ and hence

$$\frac{1}{n} \log |\mathcal{M}| \stackrel{(60)}{\geq} I(\hat{\mathbf{X}}; \mathbf{U}) = I(\hat{\mathbf{X}}; \mathbf{U}) + I(\mathbf{X}; \mathbf{U}|\hat{\mathbf{X}}) \quad (66)$$

$$= I(\mathbf{X}\hat{\mathbf{X}}; \mathbf{U}) = I(\mathbf{X}; \mathbf{U}), \quad (67)$$

where we used Prop. (iii) of Lem. 8 in (67). We also have

$$I(\mathbf{Y}; \mathbf{U}) \stackrel{(60)}{\geq} \frac{1}{n} I(\mathbf{Y}; g(\hat{\mathbf{X}})) \quad (68)$$

$$= \frac{1}{n} (H(\mathbf{Y}) + H(g(\hat{\mathbf{X}})) - H(\mathbf{Y}g(\hat{\mathbf{X}}))) \quad (69)$$

$$\geq \frac{1}{n} I(\mathbf{Y}; f(\mathbf{X})) + \frac{1}{n} |\mathcal{S}_Y|^n |\mathcal{M}|^3 \delta \log(|\mathcal{M}|^2 \delta) \\ + \frac{1}{n} |\mathcal{S}_Y|^n |\mathcal{M}|^3 \delta \log(|\mathcal{S}_Y|^n |\mathcal{M}|^2 \delta) \quad (70)$$

$$\geq \frac{1}{n} I(\mathbf{Y}; f(\mathbf{X})) + \frac{2}{n} |\mathcal{S}_Y|^n |\mathcal{M}|^3 \delta \log(|\mathcal{M}|^2 \delta) \quad (71)$$

where we used Lem. 10 in (70). Select δ such that $\varepsilon \geq -\frac{2}{n} |\mathcal{S}_Y|^n |\mathcal{M}|^3 \delta \log(|\mathcal{M}|^2 \delta)$. ■

We can now finish the proof by showing the following lemma.

Lemma 17. $\mathcal{R} \subseteq \overline{\mathcal{R}_{\text{IB}}}$.

Proof. Assume $(S, R) \in \mathcal{R}$ and choose $n \in \mathbb{N}$ and f , satisfying $\frac{1}{n} \log |\mathcal{M}| \leq R$ and (2). Choose any $\varepsilon > 0$ and find a_Y such that

$$I(a_Y^n(\mathbf{Y}); f(\mathbf{X})) \geq I(\mathbf{Y}; f(\mathbf{X})) - \varepsilon \stackrel{(2)}{\geq} nS - \varepsilon. \quad (72)$$

This is possible by applying [17, Lem. 5.2.2] with the algebra that is generated by the rectangles (cf. the paragraph above [17, Lem. 5.5.1]). We apply Proposition 16, substituting $a_Y(\mathbf{Y}) \rightarrow \mathbf{Y}$ and $\mathbf{Y} \rightarrow \mathbf{Z}$. For arbitrary $\varepsilon > 0$, Proposition 16 provides \mathbf{U} with $\mathbf{Y}a_Y(\mathbf{Y}) \dashv\!\!\!\dashv \mathbf{X} \dashv\!\!\!\dashv \mathbf{U}$ (i. e., $\mathbf{Y} \dashv\!\!\!\dashv \mathbf{X} \dashv\!\!\!\dashv \mathbf{U}$) and

$$I(\mathbf{X}; \mathbf{U}) \leq \frac{1}{n} \log |\mathcal{M}| \leq R \quad (73)$$

$$I(\mathbf{Y}; \mathbf{U}) \geq I(a_Y(\mathbf{Y}); \mathbf{U}) \quad (74)$$

$$\stackrel{(49)}{\geq} \frac{1}{n} I(a_Y^n(\mathbf{Y}); f(\mathbf{X})) - \varepsilon \stackrel{(72)}{\geq} S - 2\varepsilon. \quad (75)$$

Hence, $(S - 2\varepsilon, R) \in \mathcal{R}_{\text{IB}}$ and consequently $(S, R) \in \overline{\mathcal{R}_{\text{IB}}}$. ■

APPENDIX

A. $\hat{\mathbf{X}}$ is $\sigma(\mathbf{X})$ -measurable

For $u \in \mathcal{S}_U$ consider the $\sigma(\mathbf{X})$ -measurable function $h_u := \kappa_{U|\mathbf{X}}(\cdot; u)$ on $[0, 1]$. We obtain the vector valued function $h := (h_u)_{u \in \mathcal{S}_U}$ on $[0, 1]^{|\mathcal{S}_U|}$. This function h is $\sigma(\mathbf{X})$ -measurable as every component is $\sigma(\mathbf{X})$ -measurable. Thus, we have $\hat{\mathbf{X}}^{-1}(i) = h^{-1}(P_i) \in \sigma(\mathbf{X})$.

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